

# Transcripts and Algebraic Distances in Time Series: Stochastic Properties and Nonparametric Dependence Tests



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# Ordinal Patterns in Time Series

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Introduction

Bandt & Pompe (2002) introduced **ordinal patterns** (OPs) as complexity measures for time series characterized by *“simplicity, extremely fast calculation, robustness, and invariance with respect to nonlinear monotonous transformations”*.

**Basic idea** in time-series case: map segments

$\mathbf{X}_t = (X_t, X_{t+1}, \dots, X_{t+m-1})$  of length  $m$  from

continuously distrib., real-valued process  $(X_t)_{t \in \mathbb{Z} = \{\dots, -1, 0, 1, \dots\}}$

onto permutations from symmetric group  $S_m$  of degree  $m$ ,

where selected  $\pi_t \in S_m = \{\pi^{[1]}, \dots, \pi^{[m!]} \}$  expresses

order among values in  $\mathbf{X}_t$  in certain way (see details below).

So original process  $(X_t)$  transformed (“discretized”) into (symbolic) OP-series  $(\pi_t)$  revealing ordinal structure of  $(X_t)$ :  
**marginal distribution** of OP-series  $(\pi_t)$  provides insights into **dynamic structure** of original process  $(X_t)$ .

Expressed as  $m!$ -dimensional probability vector  $\mathbf{p}_\pi$  (or frequency vector  $\hat{\mathbf{p}}_\pi$  in case of time series data  $(x_t)$ ), with  $k$ th component being  $p_{\pi,k} = P(\pi_t = \pi^{[k]})$ .

Analogously, required for asymptotics of  $\hat{\mathbf{p}}_\pi$ ,

$\mathbf{P}_\pi(h) = (p_{\pi;kl}(h))$  with  $p_{\pi;kl}(h) = P(\pi_t = \pi^{[k]}, \pi_{t+h} = \pi^{[l]})$ .

Different (equivalent) approaches to represent OP by permutation from  $S_m$ , see Berger et al. (2019).

We focus on **permutation representation**.

Then,  $\pi = (i_1, \dots, i_m) \in S_m$  causes ascending order of components of  $\mathbf{x} = (x_1, \dots, x_m) \in \mathbb{R}^m$ , i. e.,

$$x_{i_1} \leq x_{i_2} \leq \dots \leq x_{i_m}, \quad \text{and} \quad i_{k-1} < i_k \text{ if } x_{i_{k-1}} = x_{i_k}$$

for  $k \geq 2$ , where “ $x_k = x_l$ ” if ties within  $\mathbf{x}$ . **Here:**

Probability of ties = 0, so ties at most rarely in data.

**Example:**  $(1.2, -0.7, 3.4, 1.9) \mapsto (2, 1, 4, 3),$   
 $(1.2, -0.7, 3.4, -0.7) \mapsto (2, 4, 1, 3).$

Length  $m$  of OPs (“ $m$ -OPs” for short) chosen by user.

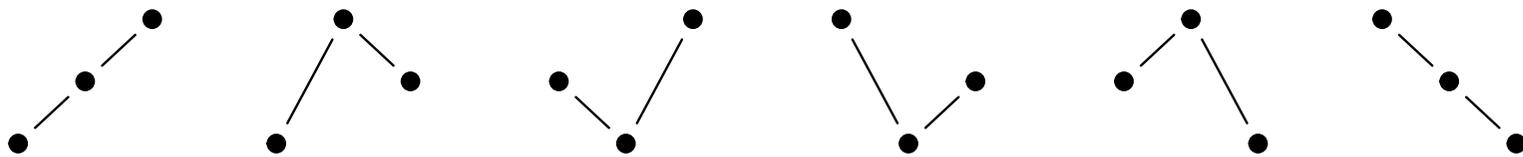
If  $m = 2$ , then downward OP (2, 1) and upward OP (1, 2), preserves only little information from original process.

However, range of  $\pi_t$  quickly increases with  $m$  as  $|S_m| = m!$ , so estimation of  $p_\pi$  quickly difficult in practice.

Therefore, in time series analysis,

**convenient choice** is  $m = 3$  (“3-OPs”; see Bandt, 2019):

$$\pi^{[1]} = (1, 2, 3) \quad (1, 3, 2) \quad (2, 1, 3) \quad (2, 3, 1) \quad (3, 1, 2) \quad (3, 2, 1) = \pi^{[6]}$$



⇒ sufficiently informative and computationally feasible.

Let  $(X_t)$  be *continuously distributed* real-valued process, independent and identically distributed (i. i. d.) *under null*.

Following **properties** crucial for dependence tests:

1. OPs invariant w.r.t. strictly monotonically increasing transformations of  $X_t$ . Thus, OPs do not depend on actual marginal distribution of  $(X_t)$  (**distribution-free** approach).
2.  $(X_t)$  is i. i. d. under null ( $\rightarrow$  exchangeability).

Thus,  $\pi_t$  discrete uniform on  $S_m$ , i. e.,  $P(\pi_t = \pi) = 1/m!$  for each  $\pi \in S_m$  (**no parameter estimation** required).

**OP-test statistics** built upon  $\hat{\mathbf{p}}_\pi$  computed from  $\pi_1, \dots, \pi_n$ , where  $\pi_t$  from  $\mathbf{X}_t = (X_t, X_{t+1}, \dots, X_{t+m-1})$  for  $t = 1, 2, \dots, n$ , and where  $\mathbf{p}_\pi = (1/m!, \dots, 1/m!)$  under i. i. d.-null:

**Theorem:** (Elsinger, 2010; Weiß, 2022)

$\sqrt{n} (\hat{\mathbf{p}}_\pi - \mathbf{p}_\pi) \rightarrow N(\mathbf{0}, \Sigma_\pi)$  with

$$\Sigma_\pi = \text{diag}(\mathbf{p}_\pi) - \mathbf{p}_\pi \mathbf{p}_\pi^\top + \sum_{h=1}^{\infty} \left( \mathbf{P}_\pi(h) + \mathbf{P}_\pi(h)^\top - 2\mathbf{p}_\pi \mathbf{p}_\pi^\top \right).$$

Distribution of OP-test statistics by “Delta method”.

**Nonparametric OP-tests**, suitable for nonlinear dependence, robust against outliers, show appealing power (Weiß, 2022).

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# Transcripts and Algebraic Distances

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Concepts & Properties

Symmetric group  $S_m$  forms **algebraic group** together with **composition** “ $\circ$ ” of permutations:

if  $\pi_1 = (i_1, i_2, \dots, i_m)$  and  $\pi_2 = (j_1, j_2, \dots, j_m)$ , then

$$\pi_1 \circ \pi_2 = (j_{i_1}, j_{i_2}, \dots, j_{i_m}).$$

**Transcript**  $\tau : S_m \times S_m \rightarrow S_m$  maps two OPs  $\pi_1, \pi_2 \in S_m$

on OP  $\pi \in S_m$  defined by

$$\pi = \tau(\pi_1, \pi_2) := \pi_2 \circ \pi_1^{-1}.$$

Hence,  $\pi = \tau(\pi_1, \pi_2)$  transforms  $\pi_1$  into  $\pi_2$ :  $\pi \circ \pi_1 = \pi_2$ .

By analogy to additive group, where  $\tau(x, y) = y - x$ ,

transcript  $\approx$  “difference”/“dissimilarity” between OPs  $\pi_1, \pi_2$ .

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## Cayley table of transcripts:

Transcript $\tau$	$\pi^{[1]}$	$\pi^{[2]}$	$\pi^{[3]}$	$\pi^{[4]}$	$\pi^{[5]}$	$\pi^{[6]}$
$\pi^{[1]} = (1, 2, 3)$	$\pi^{[1]}$	$\pi^{[2]}$	$\pi^{[3]}$	$\pi^{[4]}$	$\pi^{[5]}$	$\pi^{[6]}$
$\pi^{[2]} = (1, 3, 2)$	$\pi^{[2]}$	$\pi^{[1]}$	$\pi^{[5]}$	$\pi^{[6]}$	$\pi^{[3]}$	$\pi^{[4]}$
$\pi^{[3]} = (2, 1, 3)$	$\pi^{[3]}$	$\pi^{[4]}$	$\pi^{[1]}$	$\pi^{[2]}$	$\pi^{[6]}$	$\pi^{[5]}$
$\pi^{[4]} = (2, 3, 1)$	$\pi^{[5]}$	$\pi^{[6]}$	$\pi^{[2]}$	$\pi^{[1]}$	$\pi^{[4]}$	$\pi^{[3]}$
$\pi^{[5]} = (3, 1, 2)$	$\pi^{[4]}$	$\pi^{[3]}$	$\pi^{[6]}$	$\pi^{[5]}$	$\pi^{[1]}$	$\pi^{[2]}$
$\pi^{[6]} = (3, 2, 1)$	$\pi^{[6]}$	$\pi^{[5]}$	$\pi^{[4]}$	$\pi^{[3]}$	$\pi^{[2]}$	$\pi^{[1]}$

**Transcript**  $\approx$  “difference”/“dissimilarity” between OPs  $\pi_1, \pi_2$

$\Rightarrow$  natural relation to **distances** between permutations:

- **Cayley distance**  $d_C(\pi_1, \pi_2)$  = minimum number of transpositions to transform  $\pi_1$  into  $\pi_2$ ;
- **Kendall distance**  $d_K(\pi_1, \pi_2)$  = minimum number of adjacent transpositions to transform  $\pi_1$  into  $\pi_2$ .

$$\tau = \pi^{[1]} \Rightarrow d_C = 0, d_K = 0; \quad \tau = \pi^{[2]} \Rightarrow d_C = 1, d_K = 1;$$

$$\tau = \pi^{[3]} \Rightarrow d_C = 1, d_K = 1; \quad \tau = \pi^{[4]} \Rightarrow d_C = 2, d_K = 2;$$

$$\tau = \pi^{[5]} \Rightarrow d_C = 2, d_K = 2; \quad \tau = \pi^{[6]} \Rightarrow d_C = 1, d_K = 3;$$

see Amigó & Dale (2026) for these and further concepts.

In what follows, given process  $(X_t)$ ,  
we first compute OP-series  $(\pi_t)$  with length  $m = 3$ ,  
then **transcript series**  $(\tau_t)$  with  $\tau_t = \pi_{t+1} \circ \pi_t^{-1}$   
and corresponding **distance series**  $(d_{C,t})$  and  $(d_{K,t})$ .

Recall that successive 3-OPs “overlap”, as  
 $\pi_t$  refers to  $(X_t, X_{t+1}, X_{t+2})$  and  $\pi_{t+1}$  to  $(X_{t+1}, X_{t+2}, X_{t+3})$ .  
Also impossible transitions between successive 3-OPs,  
see de Sousa & Hlinka (2022).

**Example:**  $\pi^{[1]} = (1, 2, 3)$  never followed immediately (lag 1)  
by  $\pi^{[6]} = (3, 2, 1)$ , would only be possible with lag  $\geq 2$ .

Altogether, the  $m! = 6$  transcripts correspond to following **partition of  $4! = 24$  4-OPs**:

$\tau_t = \pi^{[1]}$	$\tau_t = \pi^{[2]}$	$\tau_t = \pi^{[3]}$	$\tau_t = \pi^{[4]}$	$\tau_t = \pi^{[5]}$	$\tau_t = \pi^{[6]}$
(1, 2, 3, 4)	(1, 2, 4, 3)	(2, 1, 3, 4)	(1, 4, 3, 2)	(1, 3, 2, 4)	(1, 3, 4, 2)
(4, 3, 2, 1)	(4, 3, 1, 2)	(3, 4, 2, 1)	(2, 1, 4, 3)	(1, 4, 2, 3)	(2, 4, 3, 1)
			(2, 4, 1, 3)	(2, 3, 1, 4)	(3, 1, 2, 4)
			(3, 2, 1, 4)	(2, 3, 4, 1)	(4, 2, 1, 3)
			(3, 2, 4, 1)	(3, 1, 4, 2)	
			(4, 1, 3, 2)	(3, 4, 1, 2)	
			(4, 2, 3, 1)	(4, 1, 2, 3)	

Therefore, marginal distribution of transcripts computed via marginal distribution of 4-OPs.

**Notations:**  $\mathbf{p}_\tau = (p_{\tau;1}, \dots, p_{\tau;m!})^\top$  with  $p_{\tau;k} = P(\tau_t = \pi^{[k]})$ ,  
 $\mathbf{P}_\tau(h) = (p_{\tau;kl}(h))$  with  $p_{\tau;kl}(h) = P(\tau_t = \pi^{[k]}, \tau_{t+h} = \pi^{[l]})$ .

Distribution of distance series via transformation matrices

$$\mathbf{T}_C = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 & 1 & 0 \end{pmatrix}, \quad \mathbf{T}_K = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}.$$

Namely,

$$\mathbf{p}_C = \mathbf{T}_C \mathbf{p}_\tau, \quad \mathbf{p}_K = \mathbf{T}_K \mathbf{p}_\tau,$$

and

$$\mathbf{P}_C(h) = \mathbf{T}_C \mathbf{P}_\tau(h) \mathbf{T}_C^\top, \quad \mathbf{P}_K(h) = \mathbf{T}_K \mathbf{P}_\tau(h) \mathbf{T}_K^\top.$$

**Proposition 3.1.1:** If  $(X_t)$  i. i. d.,  
then transcripts' marginal distribution

$$\mathbf{p}_\tau = \frac{1}{24} (2, 2, 2, 7, 7, 4)^\top.$$

Cayley and Kendall distances of successive OPs satisfy

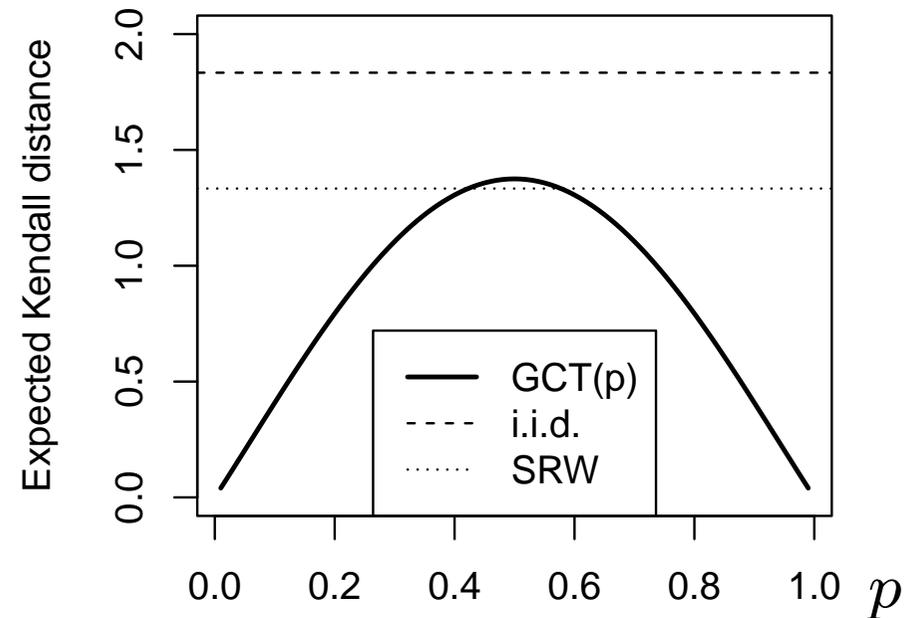
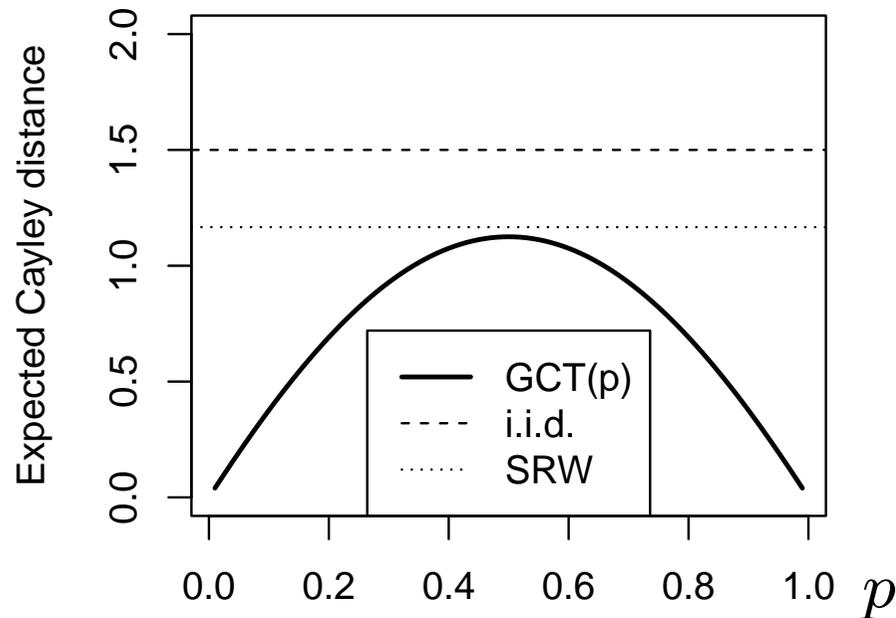
$$\mathbf{p}_C = \frac{1}{12} (1, 4, 7)^\top \quad \text{with mean } \mu_C = \frac{3}{2}, \text{ variance } \sigma_C^2 = \frac{5}{12},$$
$$\mathbf{p}_K = \frac{1}{12} (1, 2, 7, 2)^\top \quad \text{with mean } \mu_K = \frac{11}{6}, \text{ variance } \sigma_K^2 = \frac{23}{36}.$$

**Closed-form results** also for

- *symmetric random walk* of de Sousa & Hlinka (2022),
- *generalized coin-tossing* process of Silbernagel & W. (2026).

**Idea:** use transcripts/distances to test null “ $(X_t)$  is i.i.d.”.

Seems **promising approach**, for example:



Or, simulating later **alternative scenarios**,

- first-order autoregressive (AR(1)) process  $X_t = \phi X_{t-1} + \epsilon_t$  with  $\phi \in \{-0.5, 0.5\}$  and i. i. d. normal  $\epsilon_t \sim N(0, 1)$ ;
- first-order quadratic moving-average (QMA(1)) process  $X_t = \epsilon_t + 0.8 \epsilon_{t-1}^2$  with i. i. d.  $\epsilon_t \sim N(0, 1)$ ;
- first-order transposed exponential AR (TEAR(1)) process  $X_t = B_t X_{t-1} + 0.85 \epsilon_t$  with i. i. d. exponential  $\epsilon_t \sim \text{Exp}(1)$  and Bernoulli  $B_t$  with  $P(B_t = 1) = 0.15$ ;
- first-order AR conditional heteroscedasticity ARCH(1) process  $X_t = \epsilon_t \sqrt{0.2 + 0.8 X_{t-1}^2}$  with i. i. d.  $\epsilon_t \sim N(0, 1)$ ; (...)

(...) then again clear deviations from i. i. d.-case:

Process	$p_{\tau;1}$	$p_{\tau;2}$	$p_{\tau;3}$	$p_{\tau;4}$	$p_{\tau;5}$	$p_{\tau;6}$
<i>i. i. d.</i>	0.083	0.083	0.083	0.292	0.292	0.167
AR(1), +0.5	0.150	0.115	0.115	0.256	0.254	0.110
AR(1), -0.5	0.041	0.042	0.042	0.293	0.293	0.287
QMA(1)	0.103	0.076	0.094	0.348	0.220	0.160
TEAR(1)	0.101	0.085	0.089	0.233	0.341	0.152
ARCH(1)	0.100	0.085	0.085	0.279	0.279	0.172

Similar deviations for Cayley and Kendall distances.

For asymptotic implementation of dependence tests,

bivariate lag- $h$  distributions under i. i. d.-null required: (...)

Transcripts  $(\tau_t, \tau_{t+1})$  with lag 1 corresponds to occurrence of certain 5-OPs, which allow to conclude:

**3.2.1 Proposition:** If  $(X_t)$  i. i. d., then

$$\mathbf{P}_\tau(1) = \frac{1}{120} \begin{pmatrix} 2 & 1 & 1 & 3 & 3 & 0 \\ 1 & 0 & 1 & 6 & 1 & 1 \\ 1 & 1 & 0 & 1 & 6 & 1 \\ 3 & 1 & 6 & 14 & 4 & 7 \\ 3 & 6 & 1 & 4 & 14 & 7 \\ 0 & 1 & 1 & 7 & 7 & 4 \end{pmatrix},$$

$$\mathbf{P}_C(1) = \frac{1}{60} \begin{pmatrix} 1 & 1 & 3 \\ 1 & 5 & 14 \\ 3 & 14 & 18 \end{pmatrix} \quad \text{with acf } \rho_C(1) = -\frac{2}{25},$$

$$\mathbf{P}_K(1) = \frac{1}{60} \begin{pmatrix} 1 & 1 & 3 & 0 \\ 1 & 1 & 7 & 1 \\ 3 & 7 & 18 & 7 \\ 0 & 1 & 7 & 2 \end{pmatrix} \quad \text{with acf } \rho_K(1) = \frac{22}{115}.$$

Transcripts  $(\tau_t, \tau_{t+2})$  with lag 2 corresponds to occurrence of certain 6-OPs, which allow to conclude:

**3.2.3 Proposition:** If  $(X_t)$  i. i. d., then

$$\mathbf{P}_\tau(2) = \frac{1}{720} \begin{pmatrix} 2 & 7 & 7 & 18 & 18 & 8 \\ 7 & 0 & 14 & 25 & 6 & 8 \\ 7 & 14 & 0 & 6 & 25 & 8 \\ 18 & 6 & 25 & 77 & 48 & 36 \\ 18 & 25 & 6 & 48 & 77 & 36 \\ 8 & 8 & 8 & 36 & 36 & 24 \end{pmatrix},$$

$$\mathbf{P}_C(2) = \frac{1}{360} \begin{pmatrix} 1 & 11 & 18 \\ 11 & 42 & 67 \\ 18 & 67 & 125 \end{pmatrix} \quad \text{with acf } \rho_C(2) = 0,$$

$$\mathbf{P}_K(2) = \frac{1}{360} \begin{pmatrix} 1 & 7 & 18 & 4 \\ 7 & 14 & 31 & 8 \\ 18 & 31 & 125 & 36 \\ 4 & 8 & 36 & 12 \end{pmatrix} \quad \text{with acf } \rho_K(2) = \frac{8}{115}.$$

Transcripts  $(\tau_t, \tau_{t+3})$  with lag 3 corresponds to occurrence of certain 7-OPs, which allow to conclude:

**3.2.5 Proposition:** If  $(X_t)$  i. i. d., then

$$\mathbf{P}_\tau(3) = \frac{1}{5040} \begin{pmatrix} 42 & 35 & 35 & 119 & 119 & 70 \\ 35 & 40 & 30 & 116 & 129 & 70 \\ 35 & 30 & 40 & 129 & 116 & 70 \\ 119 & 129 & 116 & 421 & 440 & 245 \\ 119 & 116 & 129 & 440 & 421 & 245 \\ 70 & 70 & 70 & 245 & 245 & 140 \end{pmatrix},$$

$$\mathbf{P}_C(3) = \frac{1}{360} \begin{pmatrix} 3 & 10 & 17 \\ 10 & 40 & 70 \\ 17 & 70 & 123 \end{pmatrix} \text{ with acf } \rho_C(3) = \frac{1}{75},$$

$$\mathbf{P}_K(3) = \frac{1}{360} \begin{pmatrix} 3 & 5 & 17 & 5 \\ 5 & 10 & 35 & 10 \\ 17 & 35 & 123 & 35 \\ 5 & 10 & 35 & 10 \end{pmatrix} \text{ with acf } \rho_K(3) = \frac{1}{115}.$$

If  $m$ -OPs  $(\pi_t)$  arise from i. i. d. process  $(X_t)$ ,  
then  $m$ -OPs  $(\pi_t)$  known to be  $(m - 1)$ -dependent.

Transcripts derived from 3-OPs correspond to (sets of) 4-OPs  
 $\Rightarrow$  transcript series  $(\tau_t)$  is **3-dependent**.

(same holds for distance series  $(d_{C,t})$  and  $(d_{K,t})$ )

Therefore, sufficient to derive  $\mathbf{P}_\tau(1)$ ,  $\mathbf{P}_\tau(2)$ , and  $\mathbf{P}_\tau(3)$ .

Due to 3-dependence of  $(\tau_t)$ , it holds that

$$\mathbf{P}_\tau(h) = \mathbf{p}_\tau \mathbf{p}_\tau^\top \text{ for lag } |h| \geq 4 \quad \text{if } (X_t) \text{ is i. i. d.}$$



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# Nonparametric Tests based on Transcripts and Algebraic Distances

- Asymptotics & Performance ■

**4.1 Theorem:** Let transcript series  $(\tau_t)$  be stationary and  $\alpha$ -mixing with mixing coefficients  $\alpha_i \geq 0$ ,  $i \in \mathbb{N}_0$ , satisfying  $\sum_{i=0}^{\infty} \alpha_i < \infty$ . Then, as  $n \rightarrow \infty$ ,

$$\sqrt{n} (\hat{\mathbf{p}}_{\tau} - \mathbf{p}_{\tau}) \xrightarrow{d} \mathbf{N}(\mathbf{0}, \boldsymbol{\Sigma}_{\tau}),$$

where asymptotic covariance matrix  $\boldsymbol{\Sigma}_{\tau}$  equals

$$\boldsymbol{\Sigma}_{\tau} = \text{diag}(\mathbf{p}_{\tau}) - \mathbf{p}_{\tau} \mathbf{p}_{\tau}^{\top} + \sum_{h=1}^{\infty} \left( \mathbf{P}_{\tau}(h) + \mathbf{P}_{\tau}(h)^{\top} - 2 \mathbf{p}_{\tau} \mathbf{p}_{\tau}^{\top} \right).$$

Recall that if  $(X_t)$  is i. i. d.,

then  $(\tau_t)$  is “only” 3-dependent. Thus, (...)

**4.2 Corollary:** If  $(X_t)$  i. i. d., then  $\mathbf{p}_\tau = \frac{1}{24} (2, 2, 2, 7, 7, 4)^\top$ ,  
and as  $n \rightarrow \infty$ ,

$$\sqrt{n} (\hat{\mathbf{p}}_\tau - \mathbf{p}_\tau) \xrightarrow{d} \mathbf{N}(\mathbf{0}, \Sigma_\tau),$$

where asymptotic covariance matrix  $\Sigma_\tau$  has rank 5:

$$\Sigma_\tau = \frac{1}{20160} \begin{pmatrix} 1820 & 28 & 28 & -462 & -462 & -952 \\ 28 & 1020 & 380 & -406 & -406 & -616 \\ 28 & 380 & 1020 & -406 & -406 & -616 \\ -462 & -406 & -406 & 6259 & -4453 & -532 \\ -462 & -406 & -406 & -4453 & 6259 & -532 \\ -952 & -616 & -616 & -532 & -532 & 3248 \end{pmatrix}.$$

**4.3 Corollary:** If  $(X_t)$  i. i. d., then  $\mathbf{p}_C = \frac{1}{12} (1, 4, 7)^\top$   
and  $\mathbf{p}_K = \frac{1}{12} (1, 2, 7, 2)^\top$ , and as  $n \rightarrow \infty$ ,

$$\sqrt{n} (\hat{\mathbf{p}}_C - \mathbf{p}_C) \xrightarrow{d} N(\mathbf{0}, \Sigma_C) \quad \text{and} \quad \sqrt{n} (\hat{\mathbf{p}}_K - \mathbf{p}_K) \xrightarrow{d} N(\mathbf{0}, \Sigma_K),$$

where

$$\Sigma_C = \frac{1}{720} \begin{pmatrix} 65 & -32 & -33 \\ -32 & 128 & -96 \\ -33 & -96 & 129 \end{pmatrix} \quad \text{with rank 2,}$$

$$\Sigma_K = \frac{1}{720} \begin{pmatrix} 65 & 2 & -33 & -34 \\ 2 & 100 & -58 & -44 \\ -33 & -58 & 129 & -38 \\ -34 & -44 & -38 & 116 \end{pmatrix} \quad \text{with rank 3.}$$

**4.4 Corollary:** If  $(X_t)$  i. i. d., then  $\mu_C = \frac{3}{2}$  and  $\mu_K = \frac{11}{6}$ , and as  $n \rightarrow \infty$ , mean distances  $\bar{d}_C$  and  $\bar{d}_K$  satisfy

$$\sqrt{n}(\bar{d}_C - \mu_C) \xrightarrow{d} N(0, \frac{13}{36}) \quad \text{and} \quad \sqrt{n}(\bar{d}_K - \mu_K) \xrightarrow{d} N(0, \frac{59}{60}).$$

We obtain various **nonparametric dependence tests**:  
 $\bar{d}_C$ - **and**  $\bar{d}_K$ -**test** with (two-sided) critical or P-values,  
and upper-sided **entropy-like tests** based on

$$H(\hat{\mathbf{p}}) = - \sum_{i=1}^k \hat{p}_i (\ln \hat{p}_i - \ln p_i^{(0)}),$$

$$\Delta(\hat{\mathbf{p}}) = (\hat{\mathbf{p}} - \mathbf{p}^{(0)})^\top \text{diag}(\mathbf{p}^{(0)})^{-1} (\hat{\mathbf{p}} - \mathbf{p}^{(0)}).$$

**4.5 Proposition:** If  $(X_t)$  i. i. d., then

(a)  $-2n H_\tau(\hat{\mathbf{p}}_\tau)$  and  $n \Delta_\tau(\hat{\mathbf{p}}_\tau)$  asymptotically  $\sim \sum_{i=1}^5 \lambda_{\tau;i} \chi_{1;i}^2$  with

$$\lambda_{\tau;1} = \frac{1339}{735}, \quad \lambda_{\tau;2} \approx 1.547, \quad \lambda_{\tau;3} \approx 0.926, \quad \lambda_{\tau;4} \approx 0.718, \quad \lambda_{\tau;5} = \frac{8}{21};$$

(b)  $-2n H_C(\hat{\mathbf{p}}_C)$  and  $n \Delta_C(\hat{\mathbf{p}}_C)$  asymptotically  $\sim \sum_{i=1}^2 \lambda_{C;i} \chi_{1;i}^2$   
with

$$\lambda_{C;1} = \frac{6}{5}, \quad \lambda_{C;2} = \frac{76}{105};$$

(c)  $-2n H_K(\hat{\mathbf{p}}_K)$  and  $n \Delta_K(\hat{\mathbf{p}}_K)$  asymptotically  $\sim \sum_{i=1}^3 \lambda_{K;i} \chi_{1;i}^2$   
with

$$\lambda_{K;1} \approx 1.5468, \quad \lambda_{K;2} \approx 0.9260, \quad \lambda_{K;3} \approx 0.7177.$$

## In our paper,

- comprehensive power simulations with above alternatives;
- comparison to former OP-tests by Weiß (2022);
- illustrative data example on monthly mean air temperature (in °C) at “Hamburg–Fuhlsbüttel” (Germany) for 1936–2024.

## Quick summary:

$H_T$ - and  $\Delta_T$ -tests with appealing power for most alternatives.

For AR(1), even surpassed by  $\overline{d_K}$ -test.

$\overline{d_C}$ -test only excelled for ARCH(1). With exception of TEAR(1),

novel transcript-tests *more powerful than former OP-tests*,

and often also outperform acf.

Transcripts and algebraic distances based on OPs, stochastic properties and asymptotics of test statistics, nonparametric dependence tests with appealing power.

## **Work in progress & future research:**

- control charts based on transcripts and algebraic distances, in analogy to Weiß & Testik (2023);
- adapt to generalized OPs of Weiß & Schnurr (2024) for discrete-valued processes, or to
- spatial OPs like in Weiß & Kim (2024), or to
- cross-dependence in multiv. t. s. (Silbernagel et al., 2025).

# Thank You for Your Interest!



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