Using Spatial Ordinal Patterns for Non-parametric Testing of Spatial Dependence





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Ordinal Patterns in Time Series





Bandt & Pompe (2002) introduced **ordinal patterns** (OPs) as complexity measures for time series characterized by *'simplicity, extremely fast calculation, robustness, and invariance with respect to nonlinear monotonous transformations''*. **Basic idea** in time-series case: map segments

 $X_t = (X_t, X_{t+1}, \dots, X_{t+m-1})$ of length m from continuously distrib., real-valued process $(X_t)_{t \in \mathbb{Z} = \{\dots, -1, 0, 1, \dots\}}$ onto permutations from symmetric group S_m of order m, where selected $\pi_t \in S_m = \{\pi^{[1]}, \dots, \pi^{[m!]}\}$ expresses order among values in X_t in certain way (see details below).



In this way, original process (X_t) (or time series (x_t) thereof) transformed into (symbolic) sequence (π_t) of OPs that reveals ordinal structure of (X_t) , see Keller et al. (2007).

Marginal distribution of OP series (π_t) provides insights into dynamic structure of original process (X_t) .

Expressed as m!-dimensional probability vector p(or frequency vector \hat{p} in case of time series data (x_t)), with kth component being $p_k = P(\pi_t = \pi^{[k]})$.



Different (equivalent) approaches to represent OP by permutation from S_m , see Berger et al. (2019). We focus on **rank representation**, most intuitive approach. Then, entries of $\pi = (r_1, \ldots, r_m) \in S_m$ interpreted as ranks within $x = (x_1, \ldots, x_m) \in \mathbb{R}^m$, i. e.,

$$r_k < r_l \qquad \Leftrightarrow \qquad x_k < x_l \quad \text{or} \quad (x_k = x_l \text{ and } k < l)$$

for all $k, l \in \{1, \ldots, m\}$. Here, " $x_k = x_l$ " if ties within x.

Example: $(1.2, -0.7, 3.4, 1.9) \mapsto (2, 1, 4, 3),$ $(1.2, -0.7, 3.4, -0.7) \mapsto (3, 1, 4, 2).$



Order m of OPs chosen by user.

If m = 2, then downward OP (2,1) and upward OP (1,2),

preserves only little information from original process.

However, range of π_t quickly increases with m as $|S_m| = m!$, so estimation of p quickly difficult in practice.

Therefore, in time series analysis,

convenient choice is m = 3 (Bandt, 2019):



 \Rightarrow sufficiently informative and computationally feasible.



Let (X_t) be continuously distributed real-valued process, independent and identically distributed (i. i. d.) under null. Probability of ties = 0, so ties at most rarely in data.

Following **properties** crucial for dependence tests:

- 1. OPs invariant w.r.t. strictly monotonically increasing transformations of X_t . Thus, OPs do not depend on actual marginal distribution of $(X_t)_{\mathbb{N}}$ (**distribution-free** approach).
- 2. $(X_t)_{\mathbb{N}}$ is i. i. d. under null (\rightarrow exchangeability). Thus, π_t discrete uniform on S_m , i. e., $P(\pi_t = \pi) = 1/m!$

for each $\pi \in S_m$ (no parameter estimation required).



OP-test statistics built upon \widehat{p} computed from π_1, \ldots, π_n ,

where π_t from $X_t = (X_t, X_{t+1}, ..., X_{t+m-1})$ for t = 1, 2, ..., n.

Moving-window approach, affects asymptotics of $\sqrt{n} \left(\hat{p} - p_0 \right)$ under i. i. d.-null required, where $p_0 = (1/m!, \dots, 1/m!)$:

Theorem: (Elsinger, 2010; Weiß, 2022) $\sqrt{n} \left(\hat{p} - p_0 \right) \rightarrow \mathsf{N}(\mathbf{0}, \Sigma_m)$ with $\Sigma_m = (\sigma_{ij})_{i,j=1,...,m!}$ given by

$$\sigma_{ij} = 1/m! \left(\delta_{ij} - 1/m! \right) + \sum_{h=1}^{\infty} \left(p_{ij}(h) + p_{ji}(h) - 2/m!^2 \right).$$

Afterwards, distribution of OP-test statistics

by Taylor approximations ("Delta method").



Case m = 3:

$$\mathbf{P}(1) = \frac{1}{24} \begin{pmatrix} 1 & 1 & 0 & 0 & 2 & 0 \\ 0 & 0 & 1 & 1 & 0 & 2 \\ 1 & 2 & 0 & 0 & 1 & 0 \\ 2 & 1 & 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 2 & 0 & 1 \\ 0 & 0 & 2 & 1 & 0 & 1 \end{pmatrix}, \quad \mathbf{P}(2) = \frac{1}{120} \begin{pmatrix} 1 & 1 & 3 & 6 & 3 & 6 \\ 3 & 3 & 4 & 3 & 4 & 3 \\ 1 & 1 & 3 & 6 & 3 & 6 \\ 3 & 3 & 4 & 3 & 4 & 3 \\ 6 & 6 & 3 & 1 & 3 & 1 \\ 6 & 6 & 3 & 1 & 3 & 1 \end{pmatrix},$$

SO

$$\Sigma_3 = rac{1}{360} egin{pmatrix} 46 & -23 & -23 & 7 & 7 & -14 \ -23 & 28 & 10 & -2 & -20 & 7 \ -23 & 10 & 28 & -20 & -2 & 7 \ 7 & -23 & -20 & 28 & 10 & -23 \ 7 & -20 & -2 & 10 & 28 & -23 \ -14 & 7 & 7 & -23 & -23 & 46 \end{pmatrix}$$





Testing for Spatial Dependence in Random Fields





Real-valued and continuously distributed **spatial data** occurring in regular two-dimensional grid: $(X_t)_{t \in \mathbb{Z}^2}$. (random field, spatial process in plane, regular lattice structure) Data rectangles $(x_t) = (x_{t_1,t_2})$ with $0 \le t_1 \le m$ and $0 \le t_2 \le n$. Infer dependence in (X_t) via **spatial OPs** (SOPs), due to Ribeiro et al. (2012) and Bandt & Wittfeld (2023). $m_1 \times m_2$ -SOP computed from $m_1 \times m_2$ -rectangle from (x_t) :

- 1. concatenate rows into vector of length $m_1 \cdot m_2$,
- 2. compute corresponding $(m_1 \cdot m_2)$ th-order OP from $S_{m_1 \cdot m_2}$,
- 3. transform back into $m_1 \times m_2$ -matrix in row-wise manner.



As $|S_{m_1 \cdot m_2}| = (m_1 \cdot m_2)!$ quickly unfeasibly large, Bandt & Wittfeld (2023) recommend focus on 2 × 2-SOPs:

$$\mathbf{X}_{t} = \begin{pmatrix} X_{t_{1}-1,t_{2}-1} & X_{t_{1}-1,t_{2}} \\ X_{t_{1},t_{2}-1} & X_{t_{1},t_{2}} \end{pmatrix} =: \begin{pmatrix} x_{1} & x_{2} \\ x_{3} & x_{4} \end{pmatrix} \mapsto \begin{pmatrix} r_{1} & r_{2} \\ r_{3} & r_{4} \end{pmatrix},$$

where $(r_{1}, r_{2}, r_{3}, r_{4}) \in S_{4}$ is OP of $(x_{1}, x_{2}, x_{3}, x_{4}).$

Bandt & Wittfeld (2023): further partition into types,

$$\begin{split} &\mathcal{S}_{1} \ = \ \left\{ \left(\frac{1}{3}\frac{2}{4}\right), \left(\frac{1}{2}\frac{3}{4}\right), \left(\frac{2}{4}\frac{1}{3}\right), \left(\frac{2}{1}\frac{4}{3}\right), \left(\frac{3}{4}\frac{1}{2}\right), \left(\frac{3}{1}\frac{4}{2}\right), \left(\frac{4}{3}\frac{2}{1}\right), \left(\frac{4}{2}\frac{3}{1}\right) \right\}, \\ &\mathcal{S}_{2} \ = \ \left\{ \left(\frac{1}{4}\frac{2}{3}\right), \left(\frac{1}{2}\frac{4}{3}\right), \left(\frac{2}{3}\frac{1}{4}\right), \left(\frac{2}{1}\frac{3}{4}\right), \left(\frac{3}{4}\frac{2}{1}\right), \left(\frac{3}{2}\frac{4}{1}\right), \left(\frac{4}{3}\frac{1}{2}\right), \left(\frac{4}{1}\frac{3}{2}\right) \right\}, \\ &\mathcal{S}_{3} \ = \ \left\{ \left(\frac{1}{4}\frac{3}{2}\right), \left(\frac{1}{3}\frac{4}{2}\right), \left(\frac{2}{4}\frac{3}{1}\right), \left(\frac{2}{3}\frac{4}{1}\right), \left(\frac{3}{2}\frac{1}{4}\right), \left(\frac{3}{1}\frac{2}{4}\right), \left(\frac{4}{1}\frac{2}{3}\right), \left(\frac{4}{1}\frac{2}{3}\right) \right\}. \end{split}$$



Visual representation of types

$$\begin{split} s_{1} &= \left\{ \begin{pmatrix} 1 & 2 \\ 3 & 4 \end{pmatrix}, \begin{pmatrix} 1 & 3 \\ 2 & 4 \end{pmatrix}, \begin{pmatrix} 2 & 1 \\ 4 & 3 \end{pmatrix}, \begin{pmatrix} 2 & 4 \\ 1 & 3 \end{pmatrix}, \begin{pmatrix} 3 & 1 \\ 4 & 2 \end{pmatrix}, \begin{pmatrix} 3 & 4 \\ 1 & 2 \end{pmatrix}, \begin{pmatrix} 4 & 3 \\ 2 & 1 \end{pmatrix} \right\}, \\ s_{2} &= \left\{ \begin{pmatrix} 1 & 2 \\ 4 & 3 \end{pmatrix}, \begin{pmatrix} 1 & 4 \\ 2 & 3 \end{pmatrix}, \begin{pmatrix} 2 & 1 \\ 3 & 4 \end{pmatrix}, \begin{pmatrix} 2 & 3 \\ 1 & 4 \end{pmatrix}, \begin{pmatrix} 3 & 2 \\ 4 & 1 \end{pmatrix}, \begin{pmatrix} 3 & 4 \\ 2 & 1 \end{pmatrix}, \begin{pmatrix} 4 & 3 \\ 1 & 2 \end{pmatrix} \right\}, \\ s_{3} &= \left\{ \begin{pmatrix} 1 & 3 \\ 4 & 2 \end{pmatrix}, \begin{pmatrix} 1 & 4 \\ 3 & 2 \end{pmatrix}, \begin{pmatrix} 2 & 3 \\ 4 & 1 \end{pmatrix}, \begin{pmatrix} 2 & 4 \\ 3 & 1 \end{pmatrix}, \begin{pmatrix} 3 & 2 \\ 2 & 4 \end{pmatrix}, \begin{pmatrix} 3 & 2 \\ 1 & 4 \end{pmatrix}, \begin{pmatrix} 4 & 3 \\ 2 & 1 \end{pmatrix} \right\}, \end{split}$$

by arrows along increasing rank:

$$S_{1} = \{ \mathbf{Z}, \mathbf{M}, \mathbf{\Sigma}, \mathbf{M}, \mathbf{N}, \mathbf{X}, \mathbf{M}, \mathbf{\Sigma}, \mathbf{M}, \mathbf{\Sigma} \}, \\ S_{2} = \{ \mathbf{Z}, \mathbf{M}, \mathbf{\Sigma}, \mathbf{M}, \mathbf{M}, \mathbf{L}, \mathbf{M}, \mathbf{\Sigma} \}, \\ S_{3} = \{ \mathbf{M}, \mathbf{X}, \mathbf{X}, \mathbf{K}, \mathbf{K}, \mathbf{X}, \mathbf{X}, \mathbf{K} \}.$$



3 types more feasible in small data than 24 SOPs.

Type 1: monotonic behaviour both along rows and columns.

Type 2: uniquely increase/decrease only along one direction.

Type 3: lowest and highest ranks on either main or antidiagonal.

type = rank number which shares diagonal with rank 4.

Example: arthropods data from R-package agridat:





Asymptotics of SOP frequencies under null " (X_t) are i.i.d.":

Theorem: (Weiß & Kim, 2024) $\sqrt{mn} \left(\hat{p} - p_0 \right) \stackrel{d}{\rightarrow} N(0, \Sigma)$, where Σ equals diag $(p_0) - p_0 p_0^\top + (1 - \frac{1}{m})(1 - \frac{1}{n}) \left(\mathbf{D} + \mathbf{D}^\top + \mathbf{A} + \mathbf{A}^\top - 4 p_0 p_0^\top \right)$ $+ (1 - \frac{1}{n}) \left(\mathbf{H} + \mathbf{H}^\top - 2 p_0 p_0^\top \right) + (1 - \frac{1}{m}) \left(\mathbf{V} + \mathbf{V}^\top - 2 p_0 p_0^\top \right).$

Considers different overlaps of 2×2 -SOPs:





Closed-form expressions for $\mathbf{H}, \mathbf{V}, \mathbf{D}, \mathbf{A}$ in Weiß & Kim (2024), e.g., entries of matrix 720 $\cdot \mathbf{H}$:

| | π_1 | π_2 | π_3 | π_4 | π_5 | π_6 | π_7 | π_8 | π_9 | π_{10} | π_{11} | π_{12} | π_{13} | π_{14} | π_{15} | π_{16} | π_{17} | π_{18} | π_{19} | π_{20} | π_{21} | π_{22} | π_{23} | π_{24} |
|------------|---------|---------|---------|---------|---------|---------|---------|---------|---------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|
| π_1 | 2 | 3 | 1 | 3 | 1 | 2 | 2 | 4 | 0 | 4 | 0 | 2 | 0 | 3 | 0 | 3 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_2 | 2 | 1 | 3 | 1 | 3 | 2 | 4 | 2 | 0 | 2 | 0 | 4 | 0 | 3 | 0 | 3 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π3 | 1 | 1 | 1 | 1 | 1 | 1 | 3 | 3 | 0 | 3 | 0 | 3 | 0 | 6 | 0 | 6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_4 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3 | 0 | 3 | 0 | 4 | 0 | 2 | 0 | 2 | 4 | 2 | 3 | 1 | 3 | 1 | 2 |
| π_5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 3 | 0 | 1 | 0 | 1 | 3 | 3 | 6 | 1 | 6 | 1 | 3 |
| π_6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 2 | 0 | 2 | 0 | 2 | 2 | 4 | 3 | 3 | 3 | 3 | 4 |
| π_7 | 3 | 6 | 1 | 6 | 1 | 3 | 1 | 3 | 0 | 3 | 0 | 1 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_8 | 4 | 3 | 3 | 3 | 3 | 4 | 2 | 2 | 0 | 2 | 0 | 2 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_9 | 1 | 1 | 1 | 1 | 1 | 1 | 3 | 3 | 0 | 3 | 0 | 3 | 0 | 6 | 0 | 6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{10} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3 | 0 | 3 | 0 | 2 | 0 | 4 | 0 | 4 | 2 | 2 | 1 | 3 | 1 | 3 | 2 |
| π_{11} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 3 | 0 | 1 | 0 | 1 | 3 | 3 | 6 | 1 | 6 | 1 | 3 |
| π_{12} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | T | 0 | 1 | 0 | T | 0 | 3 | 0 | 3 | 1 | 3 | 1 | 6 | 1 | 6 | 3 |
| π_{13} | 3 | 6 1 | I | 6 1 | L | 3 | 1 | 3 | 0 | 3 | 0 | 1 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{14} | 3 | 1 | 0 | 1 2 | 0 | 3 | 3 | T | 0 | T | 0 | 3 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{15} | 2 | 5 | 1 | 5 | 1 T | 2 | 2 | 4 | 6 | 4 | 6 | 2 | 2 | 5 | 2 | 5 | 0 | 2 | 1 | 1 | 1 | 1 | 1 | 1 |
| π_{16} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | с С | 0 | с С | 0 | с С | с С | 1 | 2 | 2 | 2 | 2 | 1 |
| π_{17} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 1 | 0 | 1 1 | 0 | ∠ 1 | 0 | ∠ २ | 0 | ∠ २ | ∠ 1 | 4 2 | 5 1 | 5 | 5 1 | 5 | 4 2 |
| π_{18} | 1 | 2 2 | 2 2 | 2 2 | 2 2 | 1 | 2 | 2 | 0 | 2 | 0 | 2 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{19} | 7 2 | 1 | 6 | 1 | 6 | + 2 | ∠ २ | 2 1 | 0 | ∠ 1 | 0 | ∠ २ | 0 | 1 1 | 0 | 1 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{20} | 2 | 1 | े २ | 1 | 3 | 2 | 4 | 2 | 0 | 2 | 0 | 4 | 0 | т З | 0 | ⊥ ג | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| π_{21} | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 6 | 0 | 6 | 0 | 3 | 0 | 3 | 0 | 3 | 3 | 1 | 1 | 1 | 1 | 1 | 1 |
| π_{22} | Õ | Õ | Õ | õ | Õ | Õ | Õ | 0 | 3 | õ | 3 | õ | 4 | õ | 2 | Õ | 2 | 4 | 2 | 3 | 1 | 3 | 1 | 2 |
| π_{24} | 0 | 0 | Ő | 0 | 0 | 0 | 0 | 0 | 3 | 0 | 3 | 0 | 2 | 0 | 4 | 0 | 4 | 2 | 2 | 1 | 3 | 1 | 3 | 2 |



Above Theorem simplifies if focus on **types**:

$$\mathbf{D} - p_0 \, p_0^{ op} = \mathbf{A} - p_0 \, p_0^{ op} = \mathbf{O}$$
, and

$$\mathbf{H} = \mathbf{V} = \frac{1}{180} \begin{pmatrix} 21 & 20 & 19 \\ 20 & 21 & 19 \\ 19 & 19 & 22 \end{pmatrix}.$$

Hence,

$$\begin{split} \Sigma &= \frac{1}{9} \begin{pmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{pmatrix} + \begin{pmatrix} 1 - \frac{1}{2m} - \frac{1}{2n} \end{pmatrix} \cdot \frac{1}{45} \begin{pmatrix} 1 & 0 & -1 \\ 0 & 1 & -1 \\ -1 & -1 & 2 \end{pmatrix} \\ &\to \frac{1}{45} \begin{pmatrix} 11 & -5 & -6 \\ -5 & 11 & -6 \\ -6 & -6 & 12 \end{pmatrix} \quad \text{for } m, n \to \infty. \end{split}$$



Bandt & Wittfeld (2023) proposed following type statistics:

$$\hat{\tau} = \hat{p}_1 - 1/3$$
 and $\hat{\kappa} = \hat{p}_2 - \hat{p}_3$,
 $\tilde{\tau} = \hat{p}_3 - 1/3$ and $\tilde{\kappa} = \hat{p}_1 - \hat{p}_2$.

Asymptotics under null " (X_t) are i. i. d.": (Weiß & Kim, 2024)

Corollary: $\sqrt{mn} (\hat{\tau}, \hat{\kappa}) \stackrel{d}{\rightarrow} N(0, \Sigma')$, where Σ' equals $\Sigma' = \frac{2}{9} \begin{pmatrix} 1 & 0 \\ 0 & 3 \end{pmatrix} + (1 - \frac{1}{2m} - \frac{1}{2n}) \cdot \frac{1}{45} \begin{pmatrix} 1 & 1 \\ 1 & 5 \end{pmatrix} \approx \frac{1}{45} \begin{pmatrix} 11 & 1 \\ 1 & 35 \end{pmatrix}$ $\sqrt{mn} (\tilde{\tau}, \tilde{\kappa}) \stackrel{d}{\rightarrow} N(0, \Sigma'')$, where Σ'' equals $\Sigma'' = \frac{2}{9} \begin{pmatrix} 1 & 0 \\ 0 & 3 \end{pmatrix} + (1 - \frac{1}{2m} - \frac{1}{2n}) \cdot \frac{2}{45} \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \approx \frac{4}{45} \begin{pmatrix} 3 & 0 \\ 0 & 8 \end{pmatrix}$.





Testing for Spatial Dependence in Random Fields





Performance analysis by simulations in Weiß & Kim (2024).

Brief summary:

While spatial ACF superior for linear unilateral DGPs, e.g.,

$$X_{t_1,t_2} = \alpha_1 \cdot X_{t_1-1,t_2} + \alpha_2 \cdot X_{t_1,t_2-1} + \alpha_3 \cdot X_{t_1-1,t_2-1} + \epsilon_{t_1,t_2},$$

SOP-based tests often superior in presence of *outliers*, for *non-linear* DGPs, such as

$$X_{t_1,t_2} = \beta_1 \cdot \epsilon_{t_1-1,t_2}^2 + \beta_2 \cdot \epsilon_{t_1,t_2-1}^2 + \beta_3 \cdot \epsilon_{t_1-1,t_2-1}^2 + \epsilon_{t_1,t_2},$$
(...)



(...) and also for *bilateral* spatial DGPs, among others

$$\begin{split} X_{t_1,t_2} &= a_1 \cdot X_{t_1-1,t_2} + a_2 \cdot X_{t_1,t_2-1} + a_3 \cdot X_{t_1,t_2+1} + a_4 \cdot X_{t_1+1,t_2} + \epsilon_{t_1,t_2} \\ \text{and} \end{split}$$

 $X_{t_1,t_2} = b_1 \cdot \epsilon_{t_1-1,t_2-1}^2 + b_2 \cdot \epsilon_{t_1+1,t_2-1} + b_3 \cdot \epsilon_{t_1+1,t_2+1}^2 + b_4 \cdot \epsilon_{t_1-1,t_2+1} + \epsilon_{t_1,t_2}.$

 $\tilde{\tau}$ -test performs particularly well!

Weiß & Kim (2024): two **agricultural data examples**, e.g., **yield of barley** (in kg) in 28×7 -grid (m = 27, n = 6) from uniformity trial experiment (Kempton & Howes, 1981). Yield as deviations in 0.01 kg-units from mean yield 2.63 kg.





Except $\tilde{\kappa}$, all SOP-tests significant spatial dependence.

Type 3 too rare (7.4%), types 1 and 2 too frequent (46.9% resp. 45.7%).

Those SOPs too frequent, where maximal ranks within columns.

Explanation:

"sowing, harvesting and all intermediate farming practices were carried out column by column, and this could produce intra-column correlations" (Kempton & Howes, 1981)



OPs are well-interpretable, robust, and

flexibly adapted to different types of dependence.

If data continuously distributed, we get non-parametric tests.

Work in progress & future research:

• SOP-based hypothesis tests,

which use a refined definition of types;

- control charts based on SOPs and types,
 in analogy to Weiß & Testik (2023);
- SOPs based on "generalized OPs", where ties are explicitly accounted for, in analogy to Weiß & Schnurr (2023).

Thank You for Your Interest!





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Bandt (2019) Small order patterns ... Entropy 21, 613. Bandt & Pompe (2002) Permutation entropy ... Phys Rev L 88, 174102. Bandt & Wittfeld (2023) Two new parameters ... Chaos 33, 043124. Berger et al. (2019) Teaching ordinal patterns ... Entropy 21, 1023. Elsinger (2010) Independence tests ... Working paper 165, Ost. Nat.bank. Keller et al. (2007) Time series from ... Stoch Dyn 7, 247–272. Kempton & Howes (1981) The use of neighbouring ... JRSS-C 30, 59–70. Ribeiro et al. (2012) Complexity-entropy ... PLoS ONE 7, e40689. Weiß (2022) Non-parametric tests ... Chaos **32**, 093107. Weiß & Kim (2024) Using spatial ordinal patterns ... Spat Stat 59, 100800. Weiß & Schnurr (2023) Generalized OPs ... J Nonpar Stat, in press. Weiß & Testik (2023) Non-param. control ch... Technomet 65, 340–350.