

Risk-Based Metrics for Performance Evaluation of Control Charts



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Performance Evaluation of Control Charts

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Introduction

Performance of control chart often evaluated by **run length** L , i.e., number of observations plotted on chart until alarm.

Performance metric: mean $E[L]$, **average run length** (ARL).

If process in control, L and ARL desired “sufficiently large” (e.g., $ARL \approx 370.4$) since any alarm false signal.

In contrast, alarm for out-of-control process (= true signal) desired as early as possible, small delay in detection.

(Knoth, 2006; Montgomery, 2009)

Although most common metric, chart design only relying on ARL also criticized (Kenett & Pollak, 2012; Montgomery, 2009).

For example, if (i.i.d.) Shewhart chart, then L geometric. Geometric distribution very skewed and large std. dev., so ARL very limited impression about typical RL behaviour.

Also often large discrepancy between ARL and median.

Therefore, also different performance metrics in literature, like standard deviation of L , quantiles of L (...)

(...) approach considered here:

Evaluate chart performance using **risk metrics**.

Risk associated with possibility of undesired event (e.g., loss).

Distribution of risk helps to prepare for case when undesired event occurs (Klüppelberg et al., 2014; McNeil et al., 2015).

Most basic risk metric: **Value at Risk (VaR)**,

i.e., quantile of risk's loss distribution.

$(1 - \alpha)$ -quantile of loss distribution as threshold value, which only exceeded in worst $\alpha \cdot 100\%$ of cases.

But if undesirable event occurs, then
actual loss \geq threshold value (= VaR).

Therefore, better to average VaRs or exceedances thereof
to express “typical loss” if exceedance of VaR occurs.

Idea: Risk metrics for assessing chart performance.

If control chart evaluated for

- in-control performance, then risk of early occurrence of false alarm, so risk of low run lengths;
- out-of-control performance, then risk of delayed detection, so risk of high run lengths.

Outline:

- computation of risk metrics for run length distributions;
- application to common types of control chart;
- sketch of further application area:
risk metrics to analyze estimation uncertainty
in evaluating chart performance (Phase-I analysis).



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Risk Metrics for Run Length Distributions

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Definition & Computation

Considered risk metrics: (Göb, 2011)

- discrete α -quantile of L in in-control situation:

$$\text{VaR}_{\alpha}^{\text{ic}} := \min \{x \in \mathbb{N} \mid P(L \leq x) \geq \alpha\};$$

- discrete $(1 - \alpha)$ -quantile of L in out-of-control situation:

$$\text{VaR}_{\alpha}^{\text{ooC}} := \min \{x \in \mathbb{N} \mid P(L \leq x) \geq 1 - \alpha\};$$

we have $P(L < \text{VaR}_{\alpha}^{\text{ic}}) < \alpha$ and $P(L > \text{VaR}_{\alpha}^{\text{ooC}}) \leq \alpha$,

i.e., bound for L in worst $\alpha \cdot 100\%$ of all cases.

Considered risk metrics: (Göb, 2011)

Conditional mean run length in case of violating VaR, i.e.,

- **tail conditional expectation (TCE)** in in-control situation:

$$\text{TCE}_{\alpha}^{\text{ic}} := E[L \mid L < \text{VaR}_{\alpha}^{\text{ic}}];$$

- TCE in out-of-control situation:

$$\text{TCE}_{\alpha}^{\text{OOC}} := E[L \mid L > \text{VaR}_{\alpha}^{\text{OOC}}];$$

“typical” run length in worst $\alpha \cdot 100\%$ of all cases.

Considered risk metrics: (Göb, 2011)

Possible alternative to TCE:

- **expected shortfall (ES)** by averaging VaR:

$$ES_{\alpha}^* := \frac{1}{\alpha} \int_0^{\alpha} \text{VaR}_u^* du.$$

How to compute these measures in practice?

Use L 's **survival function** $S(x) = P(L > x)$.

Then (...)

- β -quantile of L as $\min \{x \in \mathbb{N} \mid S(x) \leq 1 - \beta\}$,
- $$\text{TCE}_{\alpha}^{\text{ic}} = \text{VaR}_{\alpha}^{\text{ic}} - \frac{\text{VaR}_{\alpha}^{\text{ic}} - \sum_{y=0}^{\text{VaR}_{\alpha}^{\text{ic}}-1} S(y)}{1 - S(\text{VaR}_{\alpha}^{\text{ic}} - 1)},$$
- $$\text{TCE}_{\alpha}^{\text{oooc}} = 1 + \text{VaR}_{\alpha}^{\text{oooc}} + \frac{E[L] - \sum_{y=0}^{\text{VaR}_{\alpha}^{\text{oooc}}} S(y)}{S(\text{VaR}_{\alpha}^{\text{oooc}})},$$
- $$\text{ES}_{\alpha}^{\text{ic}} = \text{TCE}_{\alpha}^{\text{ic}} \frac{1 - S(\text{VaR}_{\alpha}^{\text{ic}} - 1)}{\alpha} + \text{VaR}_{\alpha}^{\text{ic}} \frac{S(\text{VaR}_{\alpha}^{\text{ic}} - 1) - (1 - \alpha)}{\alpha},$$
- $$\text{ES}_{\alpha}^{\text{oooc}} = \text{TCE}_{\alpha}^{\text{oooc}} \frac{S(\text{VaR}_{\alpha}^{\text{oooc}})}{\alpha} + \text{VaR}_{\alpha}^{\text{oooc}} \frac{\alpha - S(\text{VaR}_{\alpha}^{\text{oooc}})}{\alpha},$$

also see Acerbi & Tasche (2002), Göb (2011).



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Risk Performance of Common Control Charts

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Examples

Crucial point for computation:
availability of **survival function**.

Often the case in practice:

- i.i.d. Shewhart chart, geometric distribution: formula;
- CUSUM and EWMA charts for i.i.d. Gaussian data:
R-package “spc” of Knoth (2017);
- any chart that allows to (approximately) apply
MC approach of Brook & Evans (1972).

Shewhart Charts for i. i. d. Data:

$L \sim$ shifted geometric with parameter p .

Then $S(x) = P(L > x) = (1 - p)^x$ and ARL $E[L] = 1/p$.

$$\text{VaR}_\alpha^{\text{ic}} = \left\lceil \frac{\ln(1 - \alpha)}{\ln(1 - p)} \right\rceil, \quad \text{VaR}_\alpha^{\text{oooc}} = \left\lceil \frac{\ln(\alpha)}{\ln(1 - p)} \right\rceil,$$

$$\text{TCE}_\alpha^{\text{ic}} = \text{VaR}_\alpha^{\text{ic}} - 1 + \frac{1}{p} - \frac{\text{VaR}_\alpha^{\text{ic}} - 1}{1 - (1 - p)^{\text{VaR}_\alpha^{\text{ic}} - 1}}, \quad \text{TCE}_\alpha^{\text{oooc}} = \text{VaR}_\alpha^{\text{oooc}} + \frac{1}{p}.$$

$$\text{ES}_\alpha^{\text{ic}} = \frac{1 - (1 - p)^{\text{VaR}_\alpha^{\text{ic}}}}{p\alpha} - \frac{1 - \alpha}{\alpha} \text{VaR}_\alpha^{\text{ic}},$$

$$\text{ES}_\alpha^{\text{oooc}} = \frac{(1 - p)^{\text{VaR}_\alpha^{\text{oooc}}}}{p\alpha} + \text{VaR}_\alpha^{\text{oooc}}.$$

Shewhart Charts for i. i. d. Data:

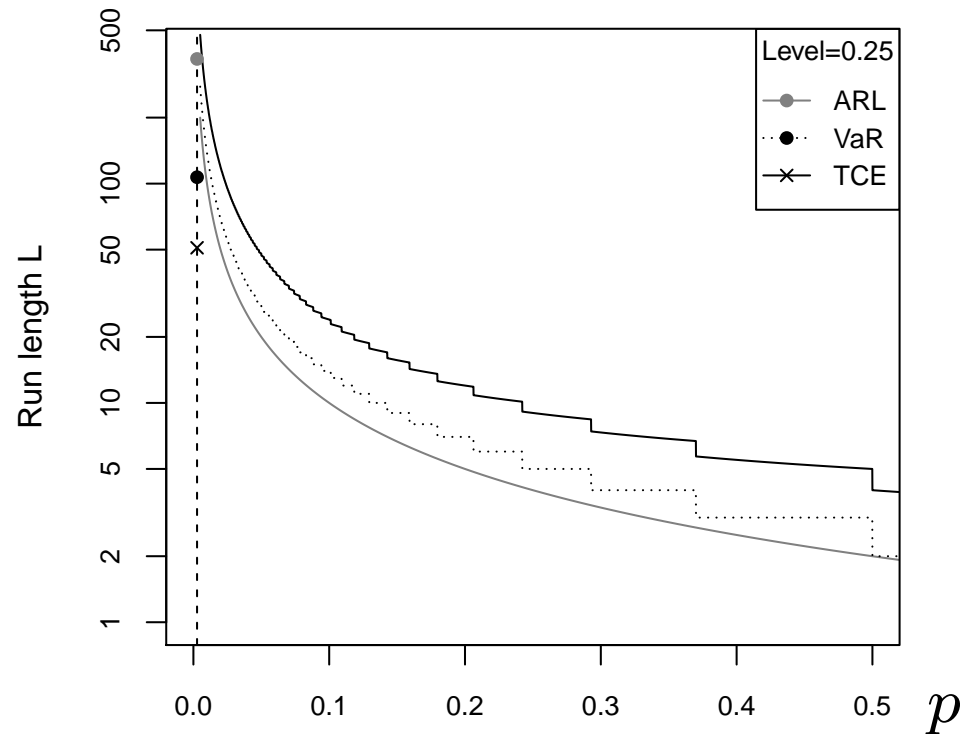
α	in-control, $ARL_0 = 370.4$					out-of-control, $ARL_1 = 10.0$				
	0.50	0.25	0.10	0.05	0.01	0.50	0.25	0.10	0.05	0.01
VaR_α	257	107	39	19	4	7	14	22	29	44
TCE_α	113.9	51.0	19.2	9.4	2.0	17.0	24.0	32.0	39.0	54.0
ES_α	114.0	51.2	19.6	9.9	2.4	16.6	23.2	31.8	38.4	53.7

Some conclusions:

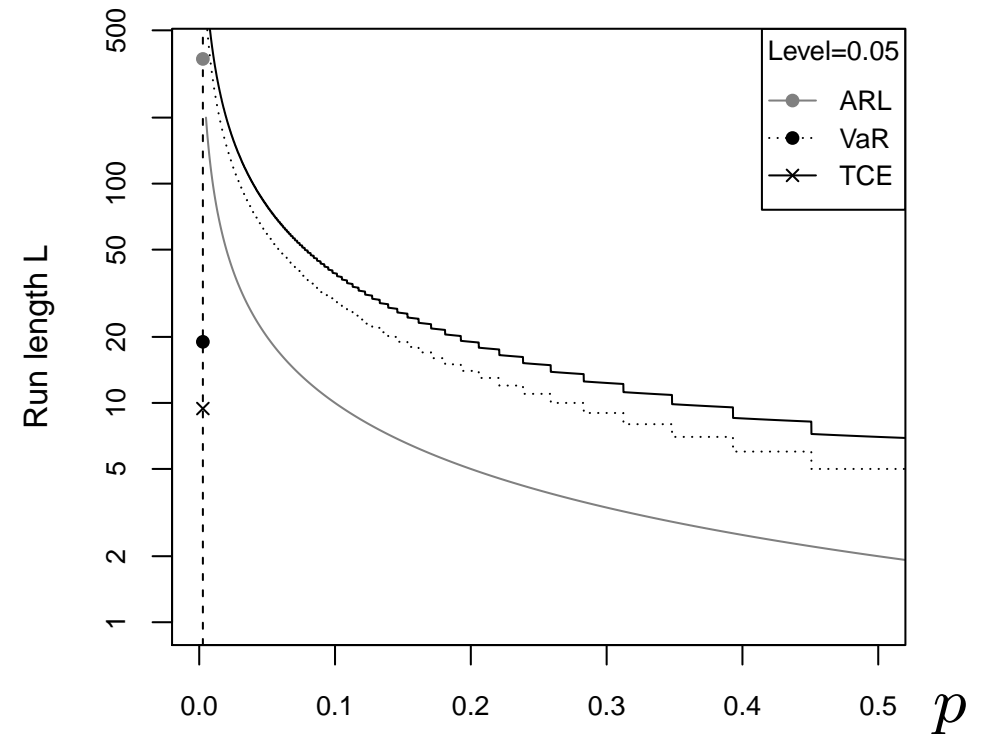
- in-control mean of L in worst 50 %
less than $1/3$ of ARL_0 ,
- out-of-control mean of L in worst 10 %
larger than 3 times ARL_1 ,
- values of TCE and ES very close to each other.

Shewhart Charts for i. i. d. Data:

ARL, VaR_α and TCE_α
for $\alpha = 0.25$:



ARL, VaR_α and TCE_α
for $\alpha = 0.05$:



CUSUM and EWMA Charts for i. i. d. Gaussian Data:

In-control VaR_α and TCE_α :

k, λ h, L	CUSUM				Shew- hart	EWMA (fixed limits)				
	0.1	0.5	1.0	2.0		0.75	0.25	0.10	0.05	0.01
	10.73	4.10	2.18	0.79		2.84	2.80	2.62	2.43	1.77
$\text{VaR}_{0.05}$	44	24	21	20	19	20	23	27	32	50
$\text{TCE}_{0.05}$	31.4	13.9	11.0	9.9	9.4	10.1	13.0	16.9	21.4	36.7
$\text{VaR}_{0.25}$	126	110	108	107	107	107	109	113	117	133
$\text{TCE}_{0.25}$	73.1	54.7	52.0	51.0	51.0	51.2	53.7	57.9	62.7	80.1
$\text{VaR}_{0.50}$	265	258	257	257	257	257	258	259	261	268
$\text{TCE}_{0.50}$	132.0	116.7	114.5	113.9	113.9	114.1	116.2	119.2	123.2	137.8

All chart designs for $\text{ARL}_0 = 370.4$,

but their in-control performances are quite different!

Two-sided EWMA Charts for i. i. d. Gaussian Data:

VaR_α and TCE_α with fixed or varying limits; $ARL_0 = 370.4$:

λ	L_{fix}	L_{var}	α	in-control, $\mu = 0$				out-of-control, $\mu = 0.5$					
				VaR_α		TCE_α		ARL		VaR_α		TCE_α	
				fix	var	fix	var	fix	var	fix	var	fix	var
0.05	2.49	2.52	0.05	32	8	21.2	3.3	26.5	21.4	56	53	70.5	67.9
			0.25	116	94	61.7	38.3			33	29	47.7	44.4
			0.50	261	251	122.7	102.8			23	18	38.1	34.4
0.10	2.70	2.71	0.05	26	14	16.1	6.0	28.2	25.7	68	66	88.0	86.3
			0.25	112	103	57.1	46.5			36	34	56.0	54.3
			0.50	259	255	118.8	110.0			23	20	43.1	40.6
0.25	2.90	2.90	0.05	22	19	12.3	9.1	41.1	40.2	114	114	151.1	151.3
			0.25	109	106	53.5	50.1			55	55	92.1	92.3
			0.50	258	257	115.9	113.3			30	29	67.1	66.3

Much lower in-control VaR and TCE values if variable limits, whereas relatively small difference in out-of-control case!

c Chart for Markov Count Processes:

MC approach applicable (exactly), analogously also for types of CUSUM and EWMA charts (bivariate MC approach), see Weiß & Testik (2009); Weiß (2018).

Here: DGP Poisson INAR(1) $X_t = \rho \circ X_{t-1} + \epsilon_t$.

Design of upper-sided c chart

if erroneously treating counts as i.i.d.,

i.e., run length performance under disregarded autocorrelation!

c Chart for Markov Count Processes:

In-control, $\mu_0 = 2.844946$ corresp. to i. i. d.- $ARL_0 = 370.4$:

ρ	ARL	VaR $_{\alpha}^{ic}$, where $\alpha =$					TCE $_{\alpha}^{ic}$, where $\alpha =$				
		0.50	0.25	0.10	0.05	0.01	0.50	0.25	0.10	0.05	0.01
0.1	372.9	259	108	40	20	4	114.7	51.4	19.7	9.9	2.0
0.2	377.9	262	109	40	20	4	116.1	51.9	19.7	9.9	2.0
0.3	386.8	268	112	41	20	4	118.7	53.3	20.1	9.9	2.0
0.4	401.5	278	116	43	21	4	123.1	55.2	21.1	10.4	2.0

Out-of-control, $\mu_1 = 5.432468$ corresp. to i. i. d.- $ARL_1 = 10.0$:

ρ	ARL	VaR $_{\alpha}^{ooc}$, where $\alpha =$					TCE $_{\alpha}^{ooc}$, where $\alpha =$				
		0.50	0.25	0.10	0.05	0.01	0.50	0.25	0.10	0.05	0.01
0.1	10.4	7	14	23	30	46	17.5	24.5	33.5	40.5	56.5
0.2	11.0	8	15	25	32	49	19.1	26.1	36.1	43.1	60.1
0.3	11.8	8	16	27	35	53	20.0	28.0	39.0	47.0	65.0
0.4	12.8	9	18	29	38	58	22.2	31.2	42.2	51.2	71.2

- Very limited impression about chart's actual performance by looking solely at ARL.
- We proposed to use risk metrics for performance evaluation as supplements to ARL metric.
- Risk metrics computed numerically exactly for many common chart types, as illustrated with number of examples.
- **Further application scenario** (see paper for details): risk metrics for evaluating control chart's ARL performance under estimation uncertainty. (...)

Phase-I analysis: estimator $\hat{\theta}_0$ of in-control parameter value θ_0 (Montgomery, 2009; Jones-Farmer et al., 2014).

Chart design for Phase-II monitoring based on $\hat{\theta}_0$, i.e., alarm triggered if

$$Z_t \notin [l(\hat{\theta}_0); u(\hat{\theta}_0)].$$

Therefore, performance metrics become function of $\hat{\theta}_0$.

E.g., $ARL(\theta | \hat{\theta}_0)$, where probability is zero that $\hat{\theta}_0$ equals θ_0 .

So $ARL(\theta_0 | \hat{\theta}_0)$ deviates from ARL_0 with certainty.

One way to quantify this deviation:

mean of conditional $\text{ARL}(\boldsymbol{\theta} | \hat{\boldsymbol{\theta}}_0)$ distribution over $\hat{\boldsymbol{\theta}}_0$,
i.e., **marginal ARL**, see Jensen et al. (2006); Testik (2007).

We propose to also compute **risk metrics for ARL**, e.g.,

$$\text{AVaR}_\alpha^{\text{ic}} := \inf \left\{ x \in (1; \infty) \mid P(\text{ARL}(\boldsymbol{\theta}_0 | \hat{\boldsymbol{\theta}}_0) \leq x) \geq \alpha \right\},$$

$$\begin{aligned} \text{ATCE}_\alpha^{\text{ic}} &:= E \left[\text{ARL}(\boldsymbol{\theta}_0 | \hat{\boldsymbol{\theta}}_0) \mid \text{ARL}(\boldsymbol{\theta}_0 | \hat{\boldsymbol{\theta}}_0) < \text{AVaR}_\alpha^{\text{ic}} \right] \\ &= \text{AVaR}_\alpha^{\text{ic}} - \frac{\text{AVaR}_\alpha^{\text{ic}} - 1 - \int_1^{\text{AVaR}_\alpha^{\text{ic}}} S_{\text{ARL}}(x) dx}{1 - S_{\text{ARL}}(\text{AVaR}_\alpha^{\text{ic}})}. \end{aligned}$$

Details and examples in paper!

Thank You for Your Interest!



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