

# Integer-valued Autoregressive Models for Counts Showing Underdispersion



HELMUT SCHMIDT  
UNIVERSITÄT

Universität der Bundeswehr Hamburg

**MATH  
STAT**

**Christian H. Weiß**

Department of Mathematics & Statistics,  
Helmut Schmidt University, Hamburg

This talk is based on the article

**Weiß, C.H. (2013):**

*Integer-valued autoregressive models for counts showing underdispersion.*

J. Appl. Statist. **40**(9), 1931–1948.

Further details and references are provided by this article.



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# Count Data Models with Underdispersion

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Approaches & Properties

Let  $X$  be count data random variable with range  $\mathbb{N}_0$ .

Denote its mean by  $\mu_X$  and its variance by  $\sigma_X^2$ .

The “normal distribution” for count data random variables:

**Poisson distribution**,  $\text{Po}(\lambda)$  with parameter  $\lambda > 0$ ,

characterized by the **equi-cumulant property**:

$$\kappa_{X,1} = \kappa_{X,2} = \dots = \lambda.$$

In particular, Poisson random variables exhibit **equidispersion**:

$$\mu_X = \sigma_X^2 (= \lambda), \quad \text{or dispersion ratio } \frac{\sigma_X^2}{\mu_X} = 1, \quad \text{respectively.}$$

Real count data rarely exhibit perfect equidispersion.

Most common deviation from Poisson model:

**overdispersion**, i. e.,

$$\text{dispersion ratio } \frac{\sigma_X^2}{\mu_X} > 1.$$

Enumerable models like negative binomial, etc.

Nearly completely neglected in the literature:

opposite phenomenon, **underdispersion**, i. e.,

$$\text{dispersion ratio } \frac{\sigma_X^2}{\mu_X} < 1.$$

## Data example 1: strike counts.

Number of outbreaks of strikes (in 4-week periods)  
in U.K. coalmining industry (1948–1959),  
see Kendall (1961).

Empirical mean and variance as 0.994 and 0.742, respectively,  
i. e., empirical variance-mean ratio  $\approx 0.75$ : underdispersion.

More details later.

## Data example 2: emergency counts.

Number of patients between  
call for examination and first treatment  
in emergency department of children's hospital  
(July 16, 2009, 10-min intervals, 08:00:00–23:59:59),  
see Weiß (2013).

Empirical mean and variance as 2.56 and 1.87, respectively,  
i. e., empirical variance-mean ratio  $\approx 0.73$ : underdispersion.

More details later.

## How to model underdispersion?

- **Generalized Poisson distribution (GP)**  
could also be used for underdispersion, but
  - necessary to truncate range  
(truncation depends on actual parameters);
  - only approximate formulae for pmf, mean, variance, etc.
- **Double Poisson distribution (DP):**  
again, essential properties known only approximately.
- **Conway-Maxwell Poisson distribution (COM):**  
again, essential properties known only approximately.



**1st Aim:** Find models for underdispersion with

- exact formulae for properties like pmf, mean, variance, etc.;
- few model parameters, preferably two model parameters (one for mean, other for dispersion).

In the sequel, detailed description of

- Good distribution;
- power-law weighted Poisson distribution (PL distribution).

Further models discussed in Weiß (2013).

**Good distribution:** (also polylogarithmic distribution)

Denote **polylogarithm** (Jonquièrè's function) as

$$\text{Li}_\nu(z) = \sum_{x=1}^{\infty} z^x \cdot x^{-\nu} \quad \text{for } |z| < 1.$$

Two-parameter **Good distribution** has pmf

$$P(X = x) = \frac{q^{x+1} \cdot (x+1)^{-\nu}}{\text{Li}_\nu(q)} \quad \text{with } 0 < q < 1, \nu \in \mathbb{R};$$

probability generating function

$$\text{pgf}(z) = \frac{1}{z} \frac{\text{Li}_\nu(qz)}{\text{Li}_\nu(q)}; \quad (\dots)$$

**Good distribution:**  $(\dots)$

moments

$$E[(X + 1)^k] = \frac{Li_{\nu-k}(q)}{Li_{\nu}(q)};$$

in particular, mean and variance as

$$E[X] = \frac{Li_{\nu-1}(q)}{Li_{\nu}(q)} - 1, \quad V[X] = \frac{Li_{\nu-2}(q)}{Li_{\nu}(q)} - \frac{Li_{\nu-1}^2(q)}{Li_{\nu}^2(q)}.$$

(Kulasekera & Tonkyn, 1992)

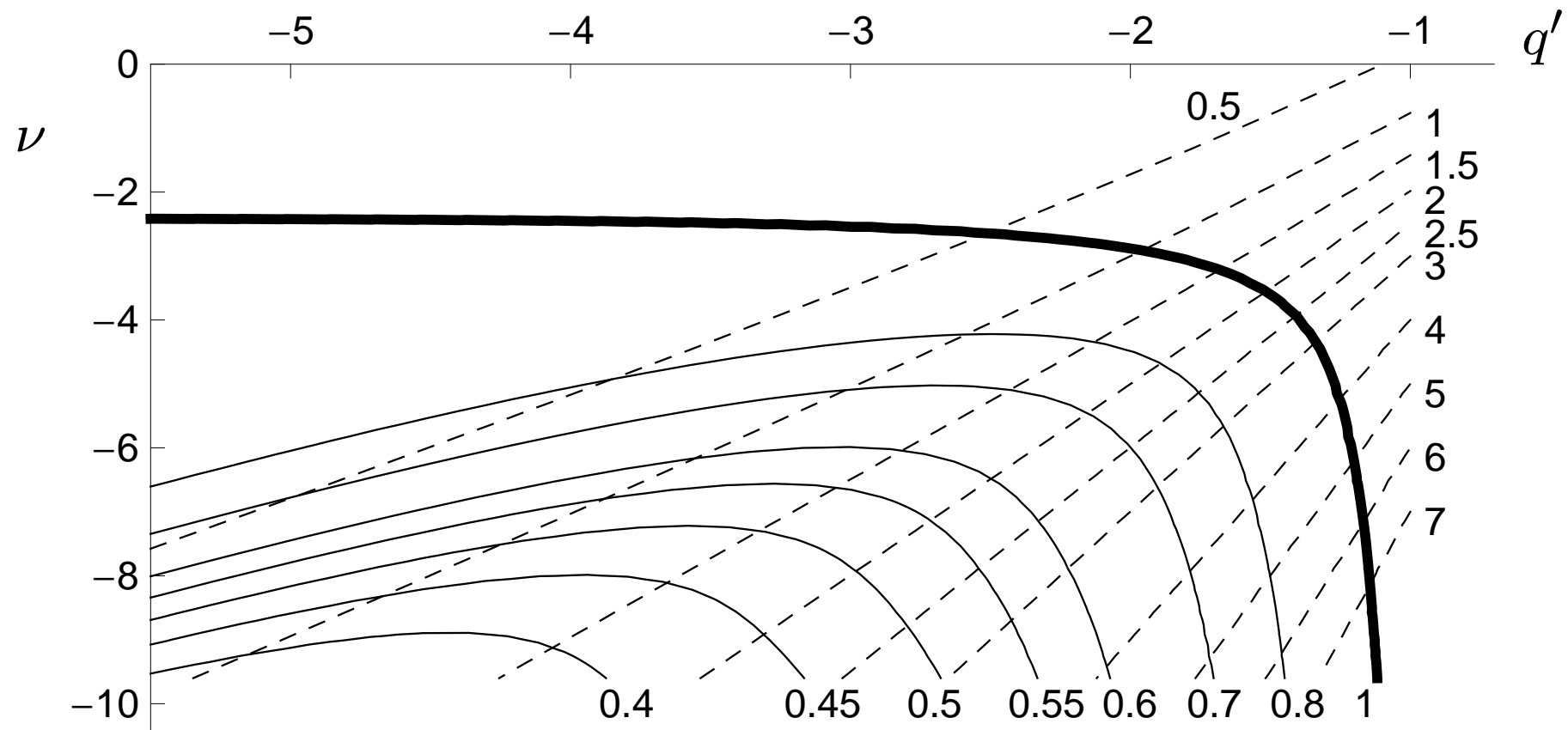
Doray & Luong (1997): **alternative parametrization,**

$q$  replaced by  $q' := \ln q \in (-\infty; 0)$ ,

advantageous in view of parameter estimation.

## Good( $q'$ , $\nu$ ) distribution:

Mean (dashed) and variance-mean ratio (solid):



**PL distribution** by del Castillo & Pérez-Casany (1998):

**Starting point:**

Let  $Y \sim \text{Po}(\lambda)$ , let  $w : \mathbb{N}_0 \rightarrow [0; \infty)$  be **weight function** such that  $0 < E_\lambda[w(Y)] < \infty$

( $w$  may depend  $\lambda$ , and on additional parameter  $\theta$ ).

Define  $X$  (**weighted version** of  $Y$ ) via pmf

$$P(X = x) = \frac{w(x)}{E_\lambda[w(Y)]} \cdot e^{-\lambda} \frac{\lambda^x}{x!} \quad \text{for } x = 0, 1, \dots$$

**PL distribution:**      (...)

Particular weight function  $w(x) = (x + a)^\nu$  with  $a > 0$  and  $\nu \in \mathbb{R}$ .

Defining

$$C(\lambda, \nu, a) = e^\lambda \cdot E_\lambda[w(Y)] = \sum_{y=0}^{\infty} \frac{\lambda^y (y + a)^\nu}{y!},$$

resulting pmf as

$$P(X = x) = \frac{\lambda^x \cdot (x + a)^\nu}{C(\lambda, \nu, a) \cdot x!};$$

$$\text{pgf}(z) = \frac{C(\lambda z, \nu, a)}{C(\lambda, \nu, a)}; \quad (\dots)$$

**PL distribution:**  $(\dots)$

factorial moments

$$E[X \cdots (X - k + 1)] = \lambda^k \cdot \frac{C(\lambda, \nu, a + k)}{C(\lambda, \nu, a)}.$$

del Castillo & Pérez-Casany (1998):

- equidispersion iff  $\nu = 0$  ( $\equiv$  Po( $\lambda$ )-distribution),
- overdispersion iff  $\nu < 0$ ,
- underdispersion iff  $\nu > 0$ .

In addition: if  $\nu \in \mathbb{N}$ , then  $C(\lambda, \nu, a)$  is  $e^\lambda$  times polynomial:

$$C(\lambda, \nu, a) = e^\lambda \cdot \sum_{k=0}^{\nu} \binom{\nu}{k} a^{\nu-k} E_\lambda[Y^k].$$

**PL distribution for underdispersion:** ( . . . )

$PL_\nu(\lambda, a)$  with given  $\nu \in \mathbb{N}$ .

Closed-form expressions for the pmf, pgf, mean and variance, and always underdispersion, see Weiß (2013) for examples.

**Alternative parametrization** by defining

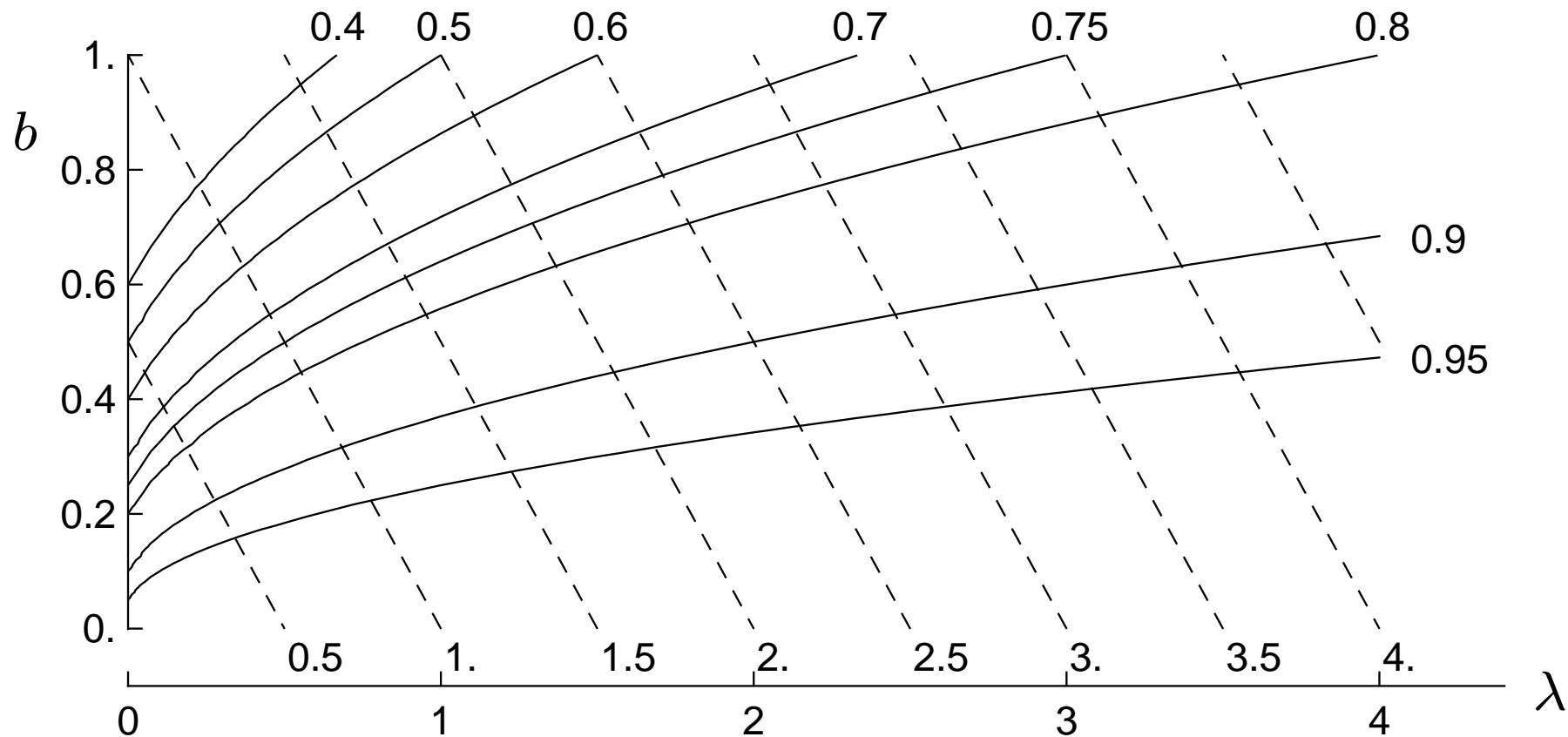
$$b := \frac{\lambda}{\lambda + a} \in (0; 1), \quad \text{equivalent to } a = \lambda \cdot \frac{1 - b}{b}.$$

$b$  more easy to interpret (Poisson corresponds to  $b \rightarrow 0$ ),  
advantageous in view of parameter estimation.



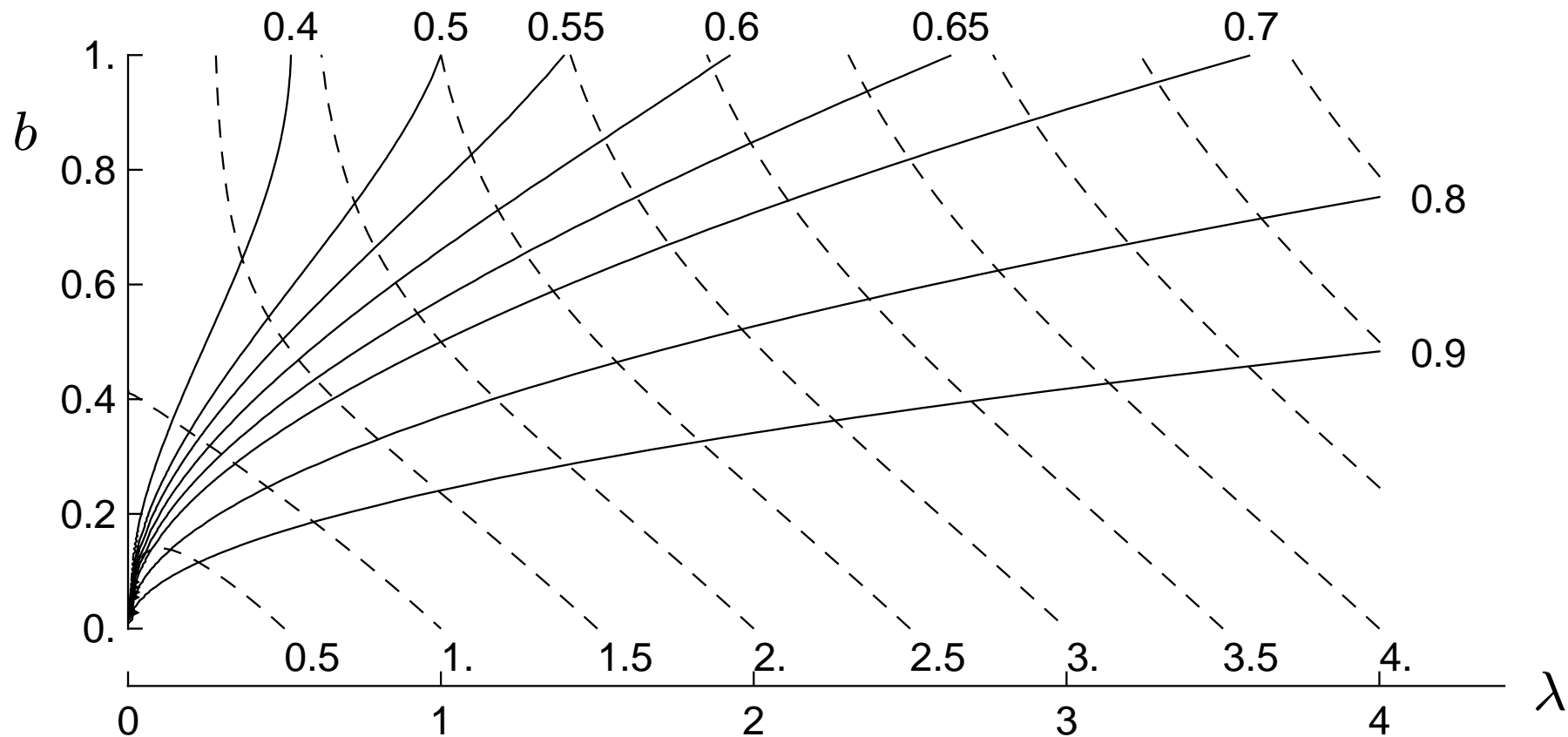
## $PL_1(\lambda, b)$ distribution:

Mean (dashed) and variance-mean ratio (solid):



## $PL_2(\lambda, b)$ distribution:

Mean (dashed) and variance-mean ratio (solid):



**Data example 1: strike counts;** mean 0.994, vm ratio 0.75.

ML estimates plus standard errors;

range of GP distribution truncated to  $\leq 7$ .

		Par. 1	Par. 2	AIC	BIC
Poisson	$(\lambda)$	0.994 (0.080)		385.87	388.92
GP	$(\alpha_0, \phi)$	0.994 (0.070)	0.873 (0.042)	381.65	387.75
Good	$(q', \nu)$	-2.866 (0.385)	-4.779 (0.751)	379.15	385.25
PL <sub>1</sub>	$(\lambda, b)$	0.467 (0.085)	0.527 (0.076)	379.20	385.30
PL <sub>2</sub>	$(\lambda, b)$	0.356 (0.081)	0.279 (0.034)	379.52	385.62



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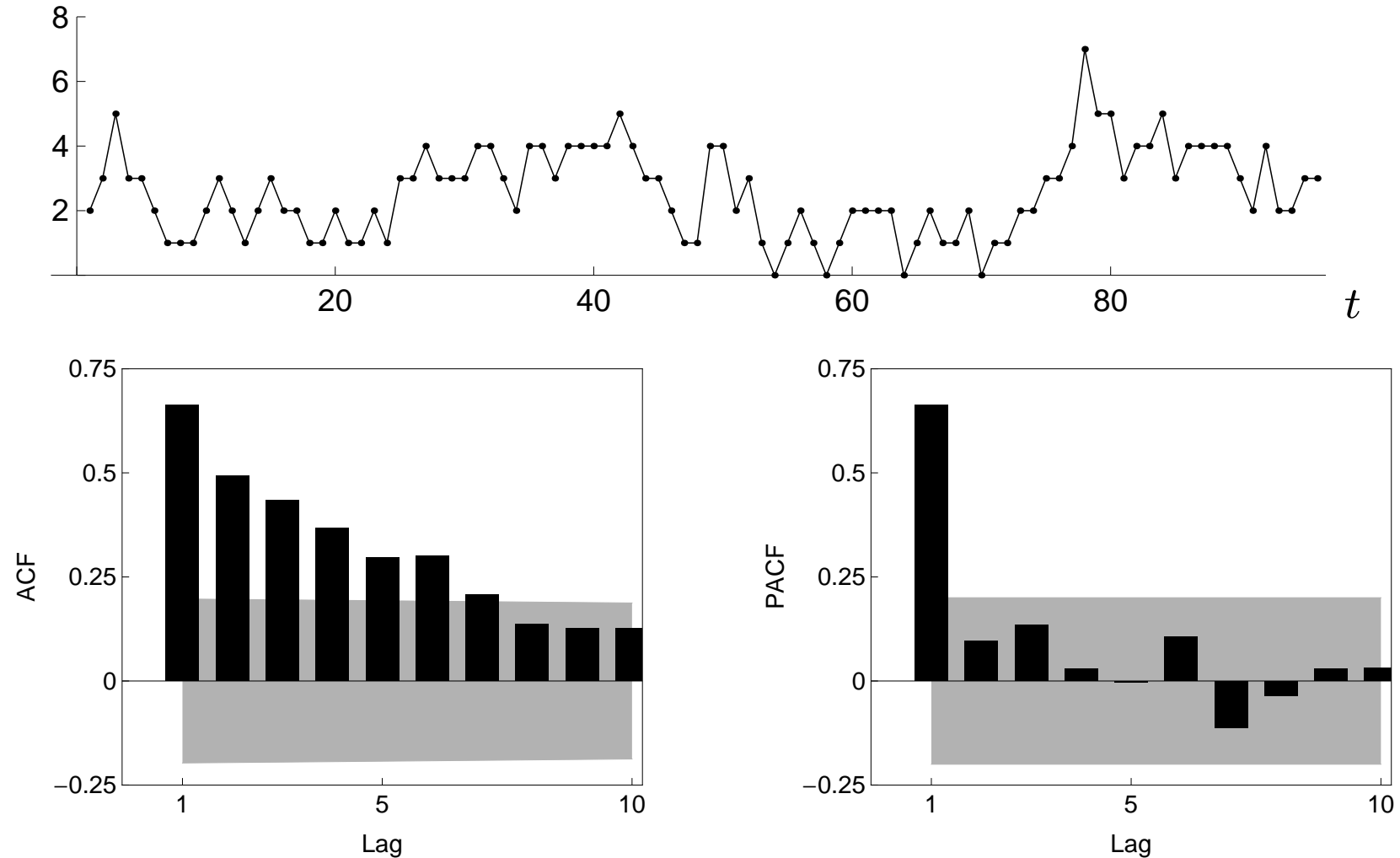
# **INAR(1) Processes with Underdispersion**

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Definition & Properties

**Data example 2: emergency counts;** mean 2.56,  $\text{vmr}$  0.73.



Popular for **real-valued** stationary processes:

**ARMA(p,q) model.** Let  $(\epsilon_t)_{\mathbb{Z}}$  white noise, then

$$X_t = \alpha_1 \cdot X_{t-1} + \dots + \alpha_p \cdot X_{t-p} + \epsilon_t + \beta_1 \cdot \epsilon_{t-1} + \dots + \beta_q \cdot \epsilon_{t-q},$$

where  $\alpha_1, \dots, \alpha_p, \beta_1, \dots, \beta_q \in \mathbb{R}$  suitably chosen.

Autocorrelation via **Yule-Walker equations:**

$$\rho_X(k) = \sum_{j=1}^p \alpha_j \cdot \rho_X(|k-j|) + \frac{\sigma_\epsilon^2}{\sigma_X^2} \cdot \sum_{i=0}^{q-k} \beta_{i+k} \cdot a_i + \delta_{k0} \cdot \frac{\sigma_\epsilon^2}{\sigma_X^2}.$$

**Example:** AR(1) model  $X_t = \alpha \cdot X_{t-1} + \epsilon_t$  with  $\rho_X(k) = \alpha^k$ .

**Not** applicable to count data processes: generally,  $\alpha \cdot X \notin \mathbb{N}_0$ .

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Several approaches in literature of  
how to avoid the “multiplication problem”.

In first part of this talk, we consider models based on  
**binomial thinning** operator (Steutel & van Harn, 1979):

$$\alpha \circ X := \sum_{i=1}^X Y_i, \quad \text{where } Y_i \text{ are i.i.d. } B(1, \alpha),$$

i. e.,  $\alpha \circ X \sim B(X, \alpha)$  and has range  $\{0, \dots, X\}$ .

( $\approx$  number of “survivors” from population of size  $X$ )

Let  $(\epsilon_t)_{\mathbb{Z}}$  be i.i.d. with range  $\mathbb{N}_0 = \{0, 1, \dots\}$ ,  
denote  $E[\epsilon_t] = \mu_\epsilon$ ,  $V[\epsilon_t] = \sigma_\epsilon^2$ . Let  $\alpha \in (0; 1)$ .

$(X_t)_{\mathbb{Z}}$  referred to as **INAR(1) process** if

$$X_t = \alpha \circ X_{t-1} + \epsilon_t,$$

plus approp. independence assumptions. (McKenzie, 1985)

## Properties:

Homogeneous Markov chain with

$$P(X_t = k \mid X_{t-1} = l) = \sum_{j=0}^{\min\{k,l\}} \binom{l}{j} \alpha^j (1 - \alpha)^{l-j} \cdot P(\epsilon_t = k - j).$$



Weiß (2013):  $\mu_\epsilon < \infty$  guarantees stationary solution.

If INAR(1) process stationary, then

$$\text{pgf}_X(z) = \text{pgf}_X(1 - \alpha + \alpha z) \cdot \text{pgf}_\epsilon(z);$$

$$\mu_X = \frac{\mu_\epsilon}{1 - \alpha}, \quad \frac{\sigma_X^2}{\mu_X} = \frac{\frac{\sigma_\epsilon^2}{\mu_\epsilon} + \alpha}{1 + \alpha}.$$

**Note:**  $X_t$  underdispersed iff  $\epsilon_t$  underdispersed.

Autocorrelation function:  $\rho_X(k) = \alpha^k$ , i. e., AR(1)-type.

For further properties and references, see Weiß (2008,2013).

## Data example 2: emergency counts;

mean 2.56, vmr 0.73,  $\hat{\rho}(1) \approx 0.664$ . We consider

- **Poisson INAR(1) model:**  $\epsilon_t \sim \text{Po}(\lambda)$ ;
- **Good INAR(1) model:**  $\epsilon_t \sim \text{Good}(q', \nu)$ ;
- **PL INAR(1) model:**  $\epsilon_t \sim \text{PL}_\nu(\lambda, b)$ ;
- **GP INAR(1) model:**  $\epsilon_t \sim \text{GP}_\nu(\alpha_0, \phi)$ ;
- **GP INARCH(1) model** by Zhu (2012).

**Note:** GP INAR(1) with innovations' range truncated  $\leq 5$ .

**Data example 2: emergency counts;** mean 2.56, vmr 0.73.

ML estimates plus standard errors.

		Par. 1	Par. 2	Par. 3	AIC	BIC
Poisson INAR(1)	$(\lambda, \alpha)$	0.737 (0.132)		0.716 (0.048)	279.84	284.97
Good INAR(1)	$(q', \nu, \alpha)$	-4.069 (1.100)	-6.887 (2.360)	0.643 (0.073)	277.48	285.18
PL <sub>1</sub> INAR(1)	$(\lambda, b, \alpha)$	0.265 (0.100)	0.687 (0.172)	0.632 (0.070)	277.04	284.73
PL <sub>2</sub> INAR(1)	$(\lambda, b, \alpha)$	0.179 (0.080)	0.318 (0.099)	0.634 (0.072)	277.35	285.04
GP INAR(1)	$(\alpha_0, \alpha, \phi)$	0.837 (0.170)	0.851 (0.091)	0.677 (0.066)	279.59	287.28
GP INARCH(1)	$(\alpha_0, \alpha_1, \phi)$	0.918 (0.190)	0.691 (0.040)	0.645 (0.078)	280.00	287.69

- Although widely neglected, underdispersion relevant phenomenon for count data random variables.
- Two-parameter Good distribution and PL distribution attractive for modeling underdispersed counts.
- INAR(1) model constitutes simple way to generate underdispersed counts with AR(1)-like dependence.
- Marginal properties derived via factorial cumulants.
- INAR(p) model by Du & Li (1992) more problematic, since underdispersed innovations do not guarantee underdispersed observations.

# Thank You for Your Interest!



HELMUT SCHMIDT  
UNIVERSITÄT

Universität der Bundeswehr Hamburg

**MATH  
STAT**

**Christian H. Weiß**

Department of Mathematics & Statistics

Helmut Schmidt University, Hamburg

[weissc@hsu-hh.de](mailto:weissc@hsu-hh.de)

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