

# Time Reversibility of INAR(1) Processes and Testing for Poisson Innovations



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# Introduction

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Motivation & Outline

Popular count-data counterpart to conventional AR(1) model:  
**INAR(1) model** by McKenzie (1985).

Let  $(\epsilon_t)_{\mathbb{Z}}$  be i.i.d. with range  $\mathbb{N}_0 = \{0, 1, \dots\}$ ,

denote  $E[\epsilon_t] = \mu_\epsilon$ ,  $V[\epsilon_t] = \sigma_\epsilon^2$ .

Let “ $\alpha \circ$ ” be **binomial thinning** operator with  $\alpha \in (0; 1)$   
(Steutel & van Harn, 1979).

$(X_t)_{\mathbb{Z}}$  referred to as **INAR(1) process** if

$$X_t = \alpha \circ X_{t-1} + \epsilon_t,$$

together with appropriate independence assumptions.

**INAR(p) models** by Alzaid & Al-Osh (1990), Du & Li (1991).

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## Properties of INAR(1) models:

Homogeneous Markov chain with

$$P(X_t = k \mid X_{t-1} = l) = \sum_{j=0}^{\min\{k,l\}} \binom{l}{j} \alpha^j (1-\alpha)^{l-j} \cdot P(\epsilon_t = k-j).$$

If INAR(1) process stationary (Heathcote, 1966), then

$$\text{pgf}_X(z) = \text{pgf}_X(1 - \alpha + \alpha z) \cdot \text{pgf}_\epsilon(z).$$

Autocorrelation function:  $\rho_X(k) = \alpha^k$ , i. e., AR(1)-type.

For further properties and references, see Weiß (2008).

Most popular instance of INAR(1) family: **Poisson** INAR(1),  
where innovations  $(\epsilon_t)_{\mathbb{Z}}$  i.i.d.  $\text{Poi}(\lambda)$ , such that  $\mu_\epsilon = \sigma_\epsilon^2 = \lambda$ .

But also many other choices for distribution of  $\epsilon_t$ , e. g.,

- with  $\sigma_\epsilon^2 > \mu_\epsilon$  (**overdispersion**) (Schweer & Weiß, 2014),
- with  $\sigma_\epsilon^2 < \mu_\epsilon$  (**underdispersion**) (Weiß, 2013).

Note that the dispersion behaviour carries over to observations:

$$\mu_X = \frac{\mu_\epsilon}{1 - \alpha}, \quad I_{\text{disp}} := \frac{\sigma_X^2}{\mu_X} = \frac{\frac{\sigma_\epsilon^2}{\mu_\epsilon} + \alpha}{1 + \alpha}.$$

How discriminate between Poisson INAR(1) (=null hypothesis)  
and alternative type of INAR(1) model?

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## Diagnostic testing of null of *Poisson INAR(1)* model:

- Meintanis & Karlis (2014): test based on empirical bivariate probability generating function; comprehensive (“omnibus test”), but computationally very demanding (complex bootstrap implementation).
- Schweer & Weiß (2014): test based on empirical dispersion index  $\hat{I}_{\text{disp}} := S_X^2 / \bar{X}$ ; easy to implement (simple closed-form formulae), powerful only if dispersion behaviour affected.

## Diagnostic testing of null of Poisson INAR(1) model:

Test procedure that is easy to implement,  
but can deal with alternatives that  
do not affect (much) dispersion behaviour?

Is there *unique feature* of Poisson INAR(1) model  
that could be used to discriminate Poisson INAR(1)  
from alternative type of INAR(1) model?

**Idea:** time-reversibility.



# Time Reversibility of INAR(1) Processes

Properties & Results

Stationary  $(X_t)_{\mathbb{Z}}$  **time-reversible** iff for all  $t_1, \dots, t_k, h \in \mathbb{Z}$ ,  
 $(X_{h+t_1}, \dots, X_{h+t_n})$  and  $(X_{h-t_1}, \dots, X_{h-t_n})$  have  
same joint distribution ( $\approx$  “symmetry in time”).

Let  $(X_t)_{\mathbb{Z}}$  be **stationary INAR(1) process**.

- Theorem 19.1 in Schweer (2015):  
Also let  $P(\epsilon = 0) \in (0; 1)$ . Then  
 $(X_t)_{\mathbb{Z}}$  **time-reversible** iff  $\epsilon$  Poisson-distributed.
- Theorem 2.2 in Schweer & Weiß (2015):  
Also let  $(X_t)_{\mathbb{Z}}$  be irreducible. Then  
 $(X_t)_{\mathbb{Z}}$  **time-reversible** iff  $\epsilon$  Poisson-distributed.

## Essentially:

Exactly *Poisson* INAR(1) processes are time-reversible.

Analogous result for continuous-valued ARMA processes:

Weiss (1975), Theorem 2:

Real-valued ARMA process is time-reversible iff  
innovations' distribution is Gaussian.

So in integer-valued case,

Poisson distribution takes place of Gaussian distribution.

Does this extend to INAR( $p$ ) models?

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Two types of INAR(p) models:

- **INAR(p) model** by Alzaid & Al-Osh (1990):
  - ✓ Theorem 19.2 in Schweer (2015): Let  $(X_t)_{\mathbb{Z}}$  be stationary AA-INAR(p) process with  $P(\epsilon = 0) \in (0; 1)$ . Then  $(X_t)_{\mathbb{Z}}$  **time-reversible** iff  $\epsilon$  Poisson-distributed.
  - ✓ Alzaid & Al-Osh (1990): AA-INAR(p) process with Poisson innovations has Poisson observations.
  - ✗ Alzaid & Al-Osh (1990): AA-INAR(p) process has ARMA(p, p-1) autocorrelation structure.

(...)

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Two types of INAR(p) models:

- **INAR(p) model** by Alzaid & Al-Osh (1990): (...)
- **INAR(p) model** by Du & Li (1991):
  - ✗ Theorem 19.3 in Schweer (2015): DL-INAR(p) process time-reversible only if parameters take degenerate values.
  - ✗ Du & Li (1991): DL-INAR(p) process with Poisson innovations does not have Poisson observations.
  - ✓ Du & Li (1991): DL-INAR(p) process has AR(p) autocorrelation structure.

Let us look back to INAR(1) case.

**Time-reversibility** is *unique feature* of Poisson INAR(1).

How can we utilize time-reversibility  
to discriminate Poisson INAR(1) model  
from alternative type of INAR(1) model?

**Idea:** Look at moments  $\mu_{i,j}(k) := E[X_t^i X_{t-k}^j]$ ;  
if time-reversible, then  $\mu_{i,j}(k) = \mu_{j,i}(k)$  has to hold.

**1<sup>st</sup> approach:** “generalized autocovariances”

$$\beta(k) := \mu_{2,1}(k) - \mu_{1,2}(k) = E[X_{t+k}^2 X_t] - E[X_{t+k} X_t^2].$$

Using moment formulae from Schweer & Weiß (2014),

$$\beta(k) = \alpha^k (1 - \alpha^k) (\bar{\mu}_{X,3} - \sigma_X^2),$$

which equals 0 for Poisson INAR(1) model.

**2<sup>nd</sup> approach:** “skewness index”

$$I_{\text{skew}} := \frac{\bar{\mu}_{X,3}}{\sigma_X^2},$$

which equals 1 for Poisson INAR(1) model.

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# Testing Time Reversibility of INAR(1) Processes

Statistics & Properties

Considered test statistics:

**1<sup>st</sup> approach:** “generalized autocovariances”

$$\hat{\beta}(k) := \frac{1}{T-k} \sum_{t=1}^{T-k} (X_{t+k}^2 X_t - X_{t+k} X_t^2),$$

which have mean *exactly* 0 if time-reversible.

**2<sup>nd</sup> approach:** “skewness index”

$$\hat{I}_{\text{skew}} := \frac{m_{X,3}}{S_X^2},$$

where  $S_X^2 = \frac{1}{T} \sum_{t=1}^T (X_t - \bar{X})^2$  and  $m_{X,3} = \frac{1}{T} \sum_{t=1}^T (X_t - \bar{X})^3$   
with  $\bar{X} = \frac{1}{T} \sum_{t=1}^T X_t$ .

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**Theorem:** Let  $(X_t)_{\mathbb{Z}}$  be Poisson INAR(1) process. Then

$$\sqrt{T-1} \cdot \hat{\beta}(1) \quad \xrightarrow{\mathcal{D}} \quad N(0, \sigma_a^2) \quad \text{for } T \rightarrow \infty,$$

where

$$\sigma_a^2 = 4\mu_X^2(1-\alpha)^2 \left( 2\frac{\alpha}{1+\alpha} + \mu_X \frac{(1-\alpha)(1+\alpha)^2}{1+\alpha+\alpha^2} \right).$$

Exact variance  $V[\sqrt{T-1}\hat{\beta}(1)]$  given by

$$\begin{aligned} \sigma_a^2 &+ \frac{2\mu_X^2}{T-1} \left( \mu_X(\mu_X + 2\alpha)^2(1 - \alpha^{T-1}) \right. \\ &\quad \left. + 2(\mu_X(1 + \alpha) + 2\alpha)^2 \frac{1 - \alpha^{2T-2}}{(1+\alpha)^2} + 6\alpha^2\mu_X \frac{1 - \alpha^{3T-3}}{(1+\alpha+\alpha^2)^2} \right). \end{aligned}$$

**Theorem:** Let  $(X_t)_{\mathbb{Z}}$  be Poisson INAR(1) process. Then

$$\sqrt{T}(\hat{I}_{\text{skew}} - 1) \xrightarrow{\mathcal{D}} N\left(0, 8 \frac{1 + \alpha^2}{1 - \alpha^2} + 6\mu_X \frac{1 + \alpha^3}{1 - \alpha^3}\right) \quad \text{for } T \rightarrow \infty.$$

Improved approximation for mean:

$$E[\hat{I}_{\text{skew}}] \approx 1 - \frac{2}{T} \frac{3 + 2\alpha + 3\alpha^2}{1 - \alpha^2}.$$

Both asymptotic results can be used to construct tests based on  $\hat{\beta}(1)$  or  $\hat{I}_{\text{skew}}$ , respectively, by plugging-in moment estimates for  $\mu_X$  and  $\alpha$ .

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## Proofs: (sketch)

- Poisson INAR(1)  $\alpha$ -mixing with expon. decreasing weights (Schweer & Weiß, 2014), existing moments of any order.
- Carries over to vector-valued process

$$\mathbf{Y}_t := (X_t - \mu_X, X_t^2 - \mu_X - \mu_X^2, X_t^3 - \mu_X - 3\mu_X^2 - \mu_X^3)^\top.$$

- CLT of Ibragimov (1962) applicable.
- Use result about cumulants for Poisson INAR(1) by Pickands & Stine (1997), Schweer & Weiß (2015):

$$\text{cum}(X_{i_1}, \dots, X_{i_r}) = \mu_X \alpha^{i_r - i_1}, \quad i_1 \leq \dots \leq i_r.$$



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# Testing Time Reversibility of INAR(1) Processes

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Empirical Investigations

**Finite-sample performance** of asymptotic approximation for mean and std. dev. of  $\hat{I}_{\text{skew}}$  (Poi. INAR(1) with  $\epsilon_t \sim \text{Poi}(\lambda)$ ):

Mean $\alpha$	$\lambda$	0.5		1		2		4	
		$T$	sim.	asympt.	$T$	sim.	asympt.	$T$	sim.
0.5	100	0.887	0.873	0.887	0.873	0.886	0.873	0.879	0.873
	250	0.950	0.949	0.948	0.949	0.951	0.949	0.953	0.949
	500	0.977	0.975	0.974	0.975	0.980	0.975	0.978	0.975
	1000	0.988	0.987	0.989	0.987	0.987	0.987	0.991	0.987

S.D. $\alpha$	$\lambda$	0.5		1		2		4	
		$T$	sim.	asympt.	$T$	sim.	asympt.	$T$	sim.
0.5	100	0.359	0.459	0.456	0.536	0.565	0.665	0.762	0.866
	250	0.265	0.290	0.312	0.339	0.394	0.420	0.523	0.548
	500	0.194	0.205	0.228	0.240	0.291	0.297	0.379	0.387
	1000	0.142	0.145	0.167	0.170	0.205	0.210	0.272	0.274

**Simulated size** for Poisson INAR(1) with  $\epsilon_t \sim \text{Poi}(\lambda)$ :

$\hat{\beta}(1)$ -test:

$\alpha$	$T$	\	$\lambda$	Size (specified parameters)				Size (estimated parameters)			
				0.5	1	2	4	0.5	1	2	4
0.5	100			0.062	0.055	0.054	0.059	0.045	0.045	0.044	0.055
	250			0.055	0.048	0.054	0.052	0.051	0.044	0.050	0.050
	500			0.052	0.050	0.049	0.053	0.053	0.050	0.050	0.053
	1000			0.056	0.049	0.051	0.052	0.054	0.049	0.050	0.050

$\hat{I}_{\text{skew}}$ -test:

$\alpha$	$T$	\	$\lambda$	Size (specified parameters)				Size (estimated parameters)			
				0.5	1	2	4	0.5	1	2	4
0.5	100			0.023	0.029	0.026	0.029	0.019	0.028	0.025	0.029
	250			0.032	0.035	0.036	0.040	0.033	0.035	0.036	0.040
	500			0.037	0.037	0.047	0.044	0.038	0.037	0.047	0.044
	1000			0.043	0.044	0.045	0.050	0.044	0.044	0.044	0.049

**Simulated power** for NB-INAR(1) with  $I_\epsilon = 1.5$ :

$\alpha$	$T$	$n$	1 ( $\mu_\epsilon = 0.5$ )			4 ( $\mu_\epsilon = 2$ )		
			$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$
0.3	100	0.556	0.370	0.394	0.600	0.283	0.289	
	250	0.869	0.542	0.723	0.915	0.399	0.557	
	500	0.987	0.738	0.938	0.997	0.540	0.827	
	1000	1.000	0.890	0.999	1.000	0.719	0.979	
0.7	100	0.226	0.328	0.126	0.234	0.129	0.085	
	250	0.435	0.646	0.276	0.440	0.244	0.164	
	500	0.677	0.884	0.460	0.697	0.523	0.268	
	1000	0.922	0.985	0.736	0.926	0.841	0.450	

$\hat{\beta}(1)$  superior for large  $\alpha$  and low  $\mu_\epsilon$ .

An **equidispersed alternative** to Poisson distribution:  
**Good distribution** (Wei , 2013).

$(q', \nu)$	Good( $q', \nu$ )			
	Mean	Var.	Skew.	Exc.
(-2.5119, -2.6470)	0.500	0.500	1.455	2.335
(-1.9560, -2.9135)	1.000	1.000	1.103	1.571
(-1.5139, -3.5533)	2.000	2.000	0.926	1.289
(-1.2498, -5.2489)	4.000	4.000	0.800	0.961

$\lambda$	Poi( $\lambda$ )			
	Mean	Var.	Skew.	Exc.
0.5	0.500	0.500	1.414	2.000
1	1.000	1.000	1.000	1.000
2	2.000	2.000	0.707	0.500
4	4.000	4.000	0.500	0.250

## Simulated power for Good-INAR(1):

$\alpha$	$T$	$(q', \nu) = (-1.5139, -3.5533)$			$(-1.2498, -5.2489)$		
		$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$
0.3	100	0.052	0.078	0.078	0.058	0.076	0.108
	250	0.066	0.084	0.126	0.060	0.091	0.174
	500	0.064	0.104	0.180	0.061	0.107	0.290
	1000	0.066	0.119	0.279	0.061	0.166	0.486
0.5	100	0.047	0.069	0.053	0.045	0.070	0.061
	250	0.054	0.086	0.078	0.053	0.081	0.104
	500	0.054	0.102	0.108	0.055	0.118	0.148
	1000	0.056	0.148	0.145	0.058	0.192	0.247

$\hat{I}_{\text{skew}}$  superior, but generally low power.

**Comparison** with  $W_{T,2}$ -test from Meintanis & Karlis (2014):  
 (Italic values for  $W_{T,2}$  taken from Table 1 in Meintanis & Karlis (2014))

Alternative model	$T = 100$				$T = 500$			
	$W_{T,2}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$	$W_{T,2}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$
NB INAR(1)								
$\mu_\epsilon = 1, \alpha = 0.5, I_\epsilon = 1.1$	<i>0.120</i>	0.074	0.078	0.046	<i>0.149</i>	0.141	0.122	0.100
$\mu_\epsilon = 1, \alpha = 0.5, I_\epsilon = 1.5$	<i>0.327</i>	0.403	0.323	0.229	<i>0.897</i>	0.941	0.782	0.718
$\mu_\epsilon = 1, \alpha = 0.5, I_\epsilon = 2.0$	<i>0.686</i>	0.798	0.651	0.538	<i>1.000</i>	1.000	0.987	0.993
Poisson INAR(2)								
$\alpha_1 = 0.5, \alpha_2 = 0.3, \lambda = 1$	<i>0.191</i>	0.353	0.116	0.100	<i>0.487</i>	0.908	0.224	0.452
$\alpha_1 = 0.5, \alpha_2 = 0.1, \lambda = 1$	<i>0.089</i>	0.082	0.060	0.039	<i>0.127</i>	0.167	0.082	0.108
$\alpha_1 = 0.3, \alpha_2 = 0.1, \lambda = 1$	<i>0.088</i>	0.055	0.047	0.035	<i>0.035</i>	0.072	0.053	0.055

$\hat{I}_{\text{disp}}$  most powerful.

**Comparison** with  $W_{T,2}$ -test from Meintanis & Karlis (2014):  
 (Italic values for  $W_{T,2}$  taken from Table 1 in Meintanis & Karlis (2014))

Alternative model	$T = 100$				$T = 500$			
	$W_{T,2}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$	$W_{T,2}$	$\hat{I}_{\text{disp}}$	$\hat{\beta}(1)$	$\hat{I}_{\text{skew}}$
INGARCH(1,1)								
$\gamma_1 = 0.4, \gamma_2 = 0.1, \delta = 0.2$	0.012	0.052	0.065	0.035	0.050	0.059	0.061	0.054
$\gamma_1 = 0.4, \gamma_2 = 0.2, \delta = 0.2$	0.042	0.082	0.078	0.059	0.108	0.172	0.101	0.145
$\gamma_1 = 0.1, \gamma_2 = 0.5, \delta = 1.0$	0.152	0.449	0.178	0.208	0.512	0.975	0.390	0.745
$\gamma_1 = 0.1, \gamma_2 = 0.5, \delta = 0.5$	0.248	0.433	0.205	0.259	0.724	0.962	0.461	0.835
$\gamma_1 = 0.4, \gamma_2 = 0.4, \delta = 0.1$	0.296	0.385	0.218	0.264	0.902	0.938	0.466	0.868
$\gamma_1 = 0.1, \gamma_2 = 0.6, \delta = 0.6$	0.376	0.686	0.273	0.374	0.988	1.000	0.608	0.967
$\gamma_1 = 0.2, \gamma_2 = 0.6, \delta = 0.1$	0.534	0.627	0.367	0.479	0.998	0.997	0.748	0.991
$\gamma_1 = 0.5, \gamma_2 = 0.45, \delta = 0.1$	0.669	0.792	0.348	0.477	0.986	1.000	0.704	0.998

$\hat{I}_{\text{disp}}$  most powerful.

- Time reversibility as unique feature of Poisson INAR(1).
  - Test statistic  $\hat{\beta}(1)$  based on generalized autocovariance as well as skewness index  $\hat{I}_{\text{skew}}$ .
  - Closed-form formulae for asymptotic distributions, hence easy to implement (no bootstrap required).
  - $\hat{\beta}(1)$ -test powerful for large autocorrelation, hence complements  $\hat{I}_{\text{disp}}$ -test.
  - $\hat{I}_{\text{skew}}$ -test useful if equidispersed alternative like Good INAR(1) model.
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# Thank You for Your Interest!



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